Total fluid consumption and risk of bladder cancer: a metaanalysis with updated data

Qinyu Liu^{1,*}, Banghua Liao^{1,*}, Ye Tian^{2,*}, Yuntian Chen¹, Deyi Luo¹, Yifei Lin¹, Hong Li¹ and Kun-Jie Wang¹

¹Urology Department, Institute of Urology (Laboratory of Reconstructive Urology), West China Hospital, Sichuan University, Chengdu, Sichuan, China

²Urology Department, Guizhou Provincial People's Hospital, Guiyang, Guizhou, China

^{*}These authors have contributed equally to this work

Correspondence to: Kun-Jie Wang, email: wangkj@scu.edu.cn

Keywords: bladder cancer, risk factor, fluid consumption, epidemiology, meta-analysis

 Received:
 July 27, 2016
 Accepted:
 May 04, 2017
 Published:
 May 23, 2017

Copyright: Liu et al. This is an open-access article distributed under the terms of the Creative Commons Attribution License 3.0 (CC BY 3.0), which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited.

ABSTRACT

With meta-analysis we tented to reveal the potential relationship between daily fluid consumption and bladder cancer risk, and to find out a recommendation on daily fluid intake. Databases of the Web of Science, PubMed and EMBASE were searched then 21 case-control and 5 cohort studies were included. Stratified analyses on gender, region, time of subjects recruiting and fluid quantity were performed as well as dose-response meta-analysis. Comparing the highest exposure category with the lowest in each study, no association appeared when all data pooled together (p=0.50), but a significant OR of 1.46 (1.02-2.08, p=0.04) was found in male subgroup. For different regions, the summarized OR was 1.44 (1.10-1.89) in American case-control studies, 1.87 (1.20-2.90) in European male subgroup and 0.24 (0.10-0.60) in Asia. There was a significant relationship that each increment 1000ml daily consumption would increase the risk by 28.6% in European male (p=0.007). Similarly every additional 1000ml consumption may increase the OR by 14.9% in American people but the association wasn't that strong (p=0.057). Stratified analyses showed fluid consumption over 3000ml/day in American residents and 2000ml/day in European male resulted in OR>1 with statistical significance. In conclusion, a relationship between higher fluid intake and higher bladder cancer risk was observed in European male and American residents and a limitation to <2000ml and <3000ml per day are recommended respectively.

INTRODUCTION

Bladder cancer is known as one of the most common malignant tumors and the risk of bladder cancer is believed to be tightly associated with life style, fluid consumption included. However, to date there still exist inconsistent conclusion on the effect of total fluid consumption on the risk. On the one hand, the urogenous contact hypothesis states that high total fluid intake can provided protective effect against bladder cancer. It believes increased fluid intake would lead to increased urine volume and urination frequency, resulting in reducing concentration of potential carcinogens in urine as well as their bladder contact time [1-3]. There are some powerful supports on this theory, including the prospective cohort study conducted by Michaud et al. which demonstrated a high fluid intake would decreased risk of bladder cancer in men [4]. On the other hand, while some source of fluid may be contaminated with carcinogens, such as chlorination byproducts and arsenic, a high-level intake may elevate the exposure of bladder cancer [5]. Apart from these, a number of recent studies found no association between total fluid intake and bladder cancer morbidity existed. To solve this controversy, a pooled analysis and a meta-analysis were conducted in 2006 and 2014 respectively [6, 7]. The former one suggested that high fluid intake may increase the risk of bladder cancer, while the latter one indicated that greater fluid consumption would reduce the bladder cancer incidence in smokers. In addition to this contradiction, both studies give no advice on recommend fluid intake volume for potential high risk or beneficial population because of lacking detailed stratified analysis on quantity.

To explored more details on this problem, we conducted a meta-analysis which added missing or updated data to the previous meta-analyses. Also we conducted more specific subgroup analyses, aiming to again assess the relationship between total fluid intake and bladder cancer risk and try to conclude recommended fluid intake volumes for different population.

RESULTS

Search result and characteristics of included studies

Using the predefined search strategies, 26 article were ultimately included in the current study [4, 8-32], including 21 case-control [8, 10, 12-20, 22-31] and five cohort studies [4, 9, 11, 21, 32] involving 12943 bladder cancer cases in total. Baseline characteristics of the included studies are presented in Supplementary (Supplementary Table 1 and Supplementary Table 2). The

flow diagram (Figure 1) showed the detailed search steps. Among these studies, ten case-control [15, 17, 22-27, 29, 30] and two cohort studies [21, 32] reported separate outcomes of male and female. Among all studies, 10 were conducted in Europe [18-23, 27, 29, 30, 32], 11 in USA [4, 8-11, 17, 24-26, 28, 31] and 5 in Asia [12-16].

Total fluid intake and risk of bladder cancer (highest vs lowest category given in each recruited study)

Risk estimates for high versus low level of daily fluid consumption are shown in Figure 2. The summary OR was 1.07 (using a random effects model, 95% CI: 0.88-1.31, P=0.50) of all studies, showing no statistically significant association between the highest fluid intake and the risk of bladder cancer. Significant heterogeneity showed among studies (P<0.001, I^2 =84%).

Total fluid intake and risk of bladder cancer by gender

The outcomes of the relationship between high-level daily fluid consumption and risk of cancer in 2 gender subgroups are shown in Figure 3. The summary OR was 1.04 (95% CI: 0.76-1.40, P=0.82) in female while was 1.46 (95% CI: 1.02-2.08, P=0.04) in male with significant association.

Total fluid intake and risk of bladder cancer by country (Supplementary Figure 1)

Subgroups could be divided into America, Europe and Asia according to the included studies. The summary



Figure 1: Flow diagram of the current meta-analysis.

OR of American case-control studies was 1.44 (95% CI: 1.10-1.89, P=0.009) with significant association. However when cohort study pooled in, the statistic significance disappeared (OR=1.22, 95% CI: 0.96-1.56, P=0.11). For European countries, there seemed to be a higher risk in people with high-level daily fluid consumption but the

association was relatively weak (OR =1.36, 95% CI: 0.99-1.87, P=0.06). And in Asia, the summary OR was 0.78 (95% CI: 0.58-1.06, P=0.11) in China and 0.09 (95% CI: 0.03-0.24, P<0.001) in Pakistan. For all studies conducted in Asia, the pooled OR was 0.24 (95% CI: 0.10-0.60, P=0.002).

Study or Subgroup	log[Odds Ratio]	SF	Weight	Odds Ratio IV, Random, 95% CI	Odds Ratio IV. Random. 95% Cl
case	iogio ano ranoj			1011100/00/	
Ahmad, 2010 [16]	-3.6929	0 7798	1.2%	0.02 [0.01, 0.11]	←
Ahmad, 2012 [12]	-2.1542		2.0%	0.12 [0.04, 0.33]	
Ahmad, 2012 [12]	-1.8245		2.1%	0.16 [0.06, 0.44]	
Baena, 2006 [19]		0.4788	2.3%	2.74 [1.07, 7.00]	
Baris, 2016 [8]	0.6206	0.211	3.8%	1.86 [1.23, 2.81]	
Bianchi, 2000 [31]	0.2788		4.4%	1.32 [1.16, 1.51]	
Bruemmer, 1997 [24]	1.5476		1.7%	4.70 [1.40, 15.78]	
Bruemmer, 1997 [24]	-0.0305		3.2%	0.97 [0.54, 1.74]	
Claude, 1986 [30]		0.3207	3.1%	4.40 [2.34, 8.24]	
Claude, 1986 [30]		1.0335	0.8%	4.00 [0.53, 30.30]	
Geoffroy-Perez, 2001 [22]	0.0677		3.8%	1.07 [0.72, 1.59]	
Geoffroy-Perez, 2001 [22]	-0.0387		2.5%	0.96 [0.42, 2.20]	
Hemelt, 2010 [15]	-0.4308		3.8%	0.65 [0.43, 0.98]	
Jensen, 1986 [29]	1.1848	0.4184	2.6%	3.27 [1.44, 7.42]	
Jensen, 1986 [29]	0.5732	0.729	1.4%	1.77 [0.43, 7.40]	
Jiang, 2008 [17]	-0.0161	0.1264	4.2%	0.98 [0.77, 1.26]	
Kunze, 1992 [27]	1.5886		2.3%	4.90 [1.95, 12.30]	
Kunze, 1992 [27]	-0.1109	0.5248	2.0%	0.90 [0.32, 2.50]	
Michaud, 2007 [18]	-0.478	0.2185	3.7%	0.62 [0.40, 0.95]	
Pohlabeln, 1999 [23]	-1.0906	0.5515	1.9%	0.34 [0.11, 0.99]	
Pohlabeln, 1999 [23]	0.4187	0.4382	2.5%	1.52 [0.64, 3.59]	
Radosavljevic, 2003 [20]	-0.0377	0.233	3.6%	0.96 [0.61, 1.52]	
Slattery, 1988 [28]	0.3075	0.2152	3.7%	1.36 [0.89, 2.07]	
Vena, 1993 [26]	1.3186	0.1943	3.9%	3.74 [2.55, 5.47]	
Wang, 2013 [10]	0.3436	0.1267	4.2%	1.41 [1.10, 1.81]	-
Wilkens, 1996 [25]	-1.2201	0.5279	2.0%	0.30 [0.10, 0.83]	
Wilkens, 1996 [25]	0.3535	0.3071	3.2%	1.42 [0.78, 2.60]	
Zhang, 2010 [14]	-0.1165	0.1603	4.0%	0.89 [0.65, 1.22]	-
Subtotal (95% CI)			80.0%	1.12 [0.88, 1.43]	•
Heterogeneity: $Tau^2 = 0.30$; Test for overall effect: $Z = 0$.		27 (P < 0	0.00001);	I² = 85%	
	02 (I = 0.00)				
cohort					
Michaud, 1999 [4]	-0.6733		3.6%	0.51 [0.32, 0.80]	
Ros, 2011 [32]	0.1133	0.133	4.2%	1.12 [0.86, 1.45]	
Zeegers, 2001 [21]		0.1757	4.0%	0.91 [0.65, 1.29]	
Zhou, 2012 [11]	0.0198		4.2%	1.02 [0.79, 1.32]	
Zhou, 2014 [9]	-0.1863	0.1546	4.1%	0.83 [0.61, 1.12]	
Subtotal (95% CI)			20.0%	0.89 [0.71, 1.11]	T
Heterogeneity: $Tau^2 = 0.04$; Test for overall effect: $Z = 1$.		P = 0.05); I² = 59%	6	
Total (95% CI)			100.0%	1.07 [0.88, 1.31]	•
Heterogeneity: $Tau^2 = 0.23$;	$Chi^2 = 203.10 df =$	32 (P < (F F F F F F F F F F F F F F F F F F F
Test for overall effect: $Z = 0$.		0- (1 (1		0470	0.05 0.2 1 5 20
Test for subgroup difference		1(P = 0)	16) $l^2 = 4$	8 2%	Lower risk Higher risk
i cor ioi subgroup unerence	5. on = 1.00, u =	· (i = 0.	10), 1 = 4	0.270	

Figure 2: Forest plot of the association between total fluid intake and bladder cancer risk (highest vs lowest quantity in each included studies). No significant association existed with all data pooled together.

а				Odds Ratio		Odds Ratio
Study or Subgroup	log[Odds Ratio]	SE	Weight	IV, Random, 95% CI	Year	IV, Random, 95% CI
case						
Jensen, 1986 [29]	0.5732	0.729	3.7%	1.77 [0.43, 7.40]	1986	
Claude, 1986 [30]	1.3855	1.0335	2.0%	4.00 [0.53, 30.30]	1986	
Kunze, 1992 [27]	-0.1109	0.5248	6.0%	0.90 [0.32, 2.50]	1992	
Wilkens, 1996 [25]	-1.2201	0.5279	6.0%	0.30 [0.10, 0.83]	1996	
Bruemmer, 1997 [24]	1.5476	0.6179	4.7%	4.70 [1.40, 15.78]	1997	
Pohlabeln, 1999 [23]	-1.0906	0.5515	5.6%	0.34 [0.11, 0.99]	1999	
Geoffroy-Perez, 2001 [22]	-0.0387	0.4228	8.0%	0.96 [0.42, 2.20]	2001	
Jiang, 2008 [17]	0.1706	0.2883	11.7%	1.19 [0.67, 2.09]	2008	- +
Hemelt, 2010 [15]	0.7839	0.4588	7.2%	2.19 [0.89, 5.38]	2010	
Subtotal (95% CI)			54.9%	1.13 [0.65, 1.96]		•
Heterogeneity: Tau ² = 0.41; 0	Chi ² = 20.92, df = 8	(P = 0.0	07); l² = 6	2%		
Test for overall effect: Z = 0.4	45 (P = 0.66)					
cohort						
Zeegers, 2001 [21]	-0.0587	0.1978	14.9%	0.94 [0.64, 1.39]	2001	
Ros, 2011 [32]	0.1398	0.2319	13.7%	1.15 [0.73, 1.81]	2011	-+
Zhou, 2014 [9]	-0.1863	0.1546	16.5%	0.83 [0.61, 1.12]	2014	
Subtotal (95% CI)			45.1%	0.93 [0.75, 1.14]		•
Heterogeneity: Tau ² = 0.00; (Chi ² = 1.38, df = 2 (P = 0.50); I² = 0%			
Test for overall effect: Z = 0.7	72 (P = 0.47)					
Total (95% CI)			100.0%	1.04 [0.76, 1.40]		•
Heterogeneity: Tau ² = 0.12; (Chi ² = 23.09, df = 1	1 (P = 0.	02); l² = 5	2%		
Test for overall effect: Z = 0.2	22 (P = 0.82)					0.1 0.2 0.5 1 2 5 10 Lower risk Higher risk
Test for subgroup differences	s: Chi² = 0.46, df =	1 (P = 0.	50), l ² = 0	%		Lower risk Flighter lisk
b				Odds Ratio		Odds Ratio

				Odds Ratio		Odds Ratio
Study or Subgroup	log[Odds Ratio]	SE	Weight	IV, Random, 95% CI	Year	IV. Random, 95% CI
case						
Claude, 1986 [30]	1.4805	0.3207	6.9%	4.40 [2.34, 8.24]	1986	
Jensen, 1986 [29]	1.1848	0.4184	6.0%	3.27 [1.44, 7.42]	1986	
Kunze, 1992 [27]	1.5886	0.4698	5.5%	4.90 [1.95, 12.30]	1992	
Vena, 1993 [26]	1.3186	0.1943	8.0%	3.74 [2.55, 5.47]	1993	
Wilkens, 1996 [25]	0.3535	0.3071	7.0%	1.42 [0.78, 2.60]	1996	
Bruemmer, 1997 [24]	-0.0305	0.2988	7.1%	0.97 [0.54, 1.74]	1997	_
Pohlabeln, 1999 [23]	0.4187	0.4382	5.8%	1.52 [0.64, 3.59]	1999	- +
Geoffroy-Perez, 2001 [22]	0.0677	0.2021	7.9%	1.07 [0.72, 1.59]	2001	
Baena, 2006 [19]	1.008	0.4788	5.5%	2.74 [1.07, 7.00]	2006	
Jiang, 2008 [17]	-0.0683	0.1435	8.4%	0.93 [0.71, 1.24]	2008	-
Hemelt, 2010 [15]	-0.5379	0.2329	7.7%	0.58 [0.37, 0.92]	2010	
Subtotal (95% Cl)			75.9%	1.77 [1.13, 2.77]		-
Heterogeneity: Tau ² = 0.47;	$Chi^2 = 76.69 df = 10$	(P < 0)	00001)· I ²	= 87%		
	om = 70.00, u = 10	J (1 . U.		0170		
		5 (i - 0.		0170		
Test for overall effect: Z = 2						
Test for overall effect: Z = 2		,	7.7%	0.51 [0.32, 0.80]	1999	
Test for overall effect: Z = 2 cohort Michaud, 1999 [4]	2.51 (P = 0.01)	0.2315	,,			- - -
Test for overall effect: Z = 2 cohort Michaud, 1999 [4] Zeegers, 2001 [21]	-0.6733	0.2315 0.1755	7.7%	0.51 [0.32, 0.80]	2001	
Test for overall effect: Z = 2 cohort Michaud, 1999 [4] Zeegers, 2001 [21] Ros, 2011 [32]	-0.6733 -0.0899	0.2315 0.1755	7.7% 8.1%	0.51 [0.32, 0.80] 0.91 [0.65, 1.29]	2001	
Test for overall effect: Z = 2 cohort Michaud, 1999 [4] Zeegers, 2001 [21] Ros, 2011 [32] Subtotal (95% CI) Heterogeneity: Tau ² = 0.09;	-0.6733 -0.0899 0.0862	0.2315 0.1755 0.1642	7.7% 8.1% 8.2% 24.1%	0.51 [0.32, 0.80] 0.91 [0.65, 1.29] 1.09 [0.79, 1.50] 0.82 [0.55, 1.23]	2001	
Test for overall effect: Z = 2 cohort Michaud, 1999 [4] Zeegers, 2001 [21] Ros, 2011 [32] Subtotal (95% CI) Heterogeneity: Tau ² = 0.09;	-0.6733 -0.0899 0.0862 Chi ² = 7.29, df = 2 (1	0.2315 0.1755 0.1642	7.7% 8.1% 8.2% 24.1%	0.51 [0.32, 0.80] 0.91 [0.65, 1.29] 1.09 [0.79, 1.50] 0.82 [0.55, 1.23]	2001	
Test for overall effect: Z = 2 cohort Michaud, 1999 [4] Zeegers, 2001 [21] Ros, 2011 [32] Subtotal (95% CI) Heterogeneity: Tau ² = 0.09; Test for overall effect: Z = 0	-0.6733 -0.0899 0.0862 Chi ² = 7.29, df = 2 (1	0.2315 0.1755 0.1642	7.7% 8.1% 8.2% 24.1%	0.51 [0.32, 0.80] 0.91 [0.65, 1.29] 1.09 [0.79, 1.50] 0.82 [0.55, 1.23]	2001	 •
Test for overall effect: Z = 2 cohort Michaud, 1999 [4] Zeegers, 2001 [21] Ros, 2011 [32] Subtotal (95% CI) Heterogeneity: Tau ² = 0.09; Test for overall effect: Z = 0 Total (95% CI)	-0.6733 -0.0899 0.0862 Chi ² = 7.29, df = 2 (f 9.95 (P = 0.34)	0.2315 0.1755 0.1642 P = 0.03	7.7% 8.1% 8.2% 24.1%); I ² = 73%	0.51 [0.32, 0.80] 0.91 [0.65, 1.29] 1.09 [0.79, 1.50] 0.82 [0.55, 1.23]	2001	
Test for overall effect: Z = 2 cohort Michaud, 1999 [4] Zeegers, 2001 [21] Ros, 2011 [32] Subtotal (95% CI)	-0.6733 -0.0899 0.0862 Chi ² = 7.29, df = 2 (1 .95 (P = 0.34) Chi ² = 98.69, df = 13	0.2315 0.1755 0.1642 P = 0.03	7.7% 8.1% 8.2% 24.1%); I ² = 73%	0.51 [0.32, 0.80] 0.91 [0.65, 1.29] 1.09 [0.79, 1.50] 0.82 [0.55, 1.23]	2001	0.1 0.2 0.5 1 2 5 10 Lower risk Higher risk

Figure 3: Forest plot of the association between total fluid intake and bladder cancer risk (highest vs lowest quantity) **in** female **(a)** and male **(b)**. Significant relationship between higher daily fluid intake and higher risk of bladder cancer only exist among males (OR=1.46, 95%CI=1.02-2.08, p=0.01).

Total fluid intake and risk of bladder cancer by recruit time

Risk estimates for the highest versus the lowest level of total fluid consumption by recruit time subgroup are shown in Supplementary Figure 2. For those studies included in this stratified analysis, 3 of 10 studies in Europe, 4 of 10 studies in America and all studies in Asia recruited participants after 1990. The summary OR of studies recruiting participants before year 1990 was 1.28 (95% CI: 1.00-1.64, P=0.05) and 0.73 (95% CI: 0.50-1.05, P=0.09) after.

The outcomes above found that significant results were only observed among case control studies, and the pooled results were different among countries and genders. So further analyses were done for possible higher risk regions.

Total fluid intake and bladder cancer risk in Europe

Total fluid intake and risk of bladder cancer in Europe by gender (Figure 4a & 4b)

As the outcomes showed, the summary OR in European female was 0.98 (95% CI: 0.73-1.32, P=0.91)

and 1.87 (95% CI: 1.20-2.90, P=0.006) in male. Significant association only existed in male.

Total fluid intake and risk of bladder cancer in European male by daily fluid intake quantity level

Since positive pooled outcome showed up in casecontrol studies among European males, we did a stratified analysis by daily fluid intake quantity level and doseresponse meta-analysis for this population (5 studies with 1973 cases and 2386 controls). According to the included studies, daily quantity (except the reference quantity) could be divide in to the following levels: the highest quantity (>3000ml), the 2nd highest quantity (2000-3000ml) and the 3rd highest quantity (<2000ml). The summary OR of these three subgroups were 2.49 (95% CI: 1.69-3.68, P<0.001), 1.67 (95% CI: 1.12-2.51, P=0.01) and 1.14 (95% CI: 0.88-1.46, P=0.32) respectively. Significant association existed when daily quantity reached over 2000ml (Figure 4c, 4d & 4e).

The dose-response meta-analysis included 4 researches (1633 cases and 2046 controls) and showed that association between cancer risk and fluid consumption matched a liner relation (p=0.007). Each increment 1000ml daily consumption would increase the risk by 28.6% (OR=1.286, 95% CI=1.071-1.544). No evidences of non-linear relationship existed. (p=0.438) (Figure 6a).



Figure 4: Forest plot of the association between total fluid intake and bladder cancer risk in Europe. (a) & (b) show the highest vs lowest quantity outcomes in European female and male, significant relationship between higher daily fluid intake and higher risk of bladder cancer only exist among males (OR=1.87, 95%CI=1.20-2.90, p<0.001). The others figures show outcomes in European male by daily fluid intake quantity level ((c) >around 3000ml, (d) around 2000-3000ml, (e) <around 2000ml). Significant association show up when the level reach over 2000ml.

Total fluid intake and bladder cancer risk in America

Total fluid intake and risk of bladder cancer in America by gender (Figure 5a & 5b)

The summary OR was 1.14 (95% CI: 0.32-4.15, P=0.84) in American female and 0.83 (95% CI: 0.61-1.12,

P=0.23) in America male. No statistic significance existed in both gender.

Total fluid intake and risk of bladder cancer in America by daily fluid intake quantity level

Since American case-control studies showed higher risk in high fluid intake population, we did a stratified analysis by daily fluid intake quantity level using the data



Figure 5: Forest plot of the association between total fluid intake and bladder cancer risk in America. (a) & (b) show the highest vs lowest quantity outcomes in European female and male and (c) shows outcomes in American residents by daily fluid intake quantity level (>3000ml, 2500-3000ml and <2500ml). Significant association show up when the level reach over 3000ml with the data from both genders.



Figure 6: Dose-response meta-analysis outcomes of data from European male and American people of both gender. There is a strong relationship in European males that each increment 1000ml daily consumption would increase the risk by 28.6% (p=0.007) (a), and a relatively weaker association in American people that every additional 1000ml daily consumption may increase the OR by 14.9% (p=0.057) (b).

of both genders from case control studies (6 studies with 4636 cases and 6274 controls). Since upper and lower boundaries in each category varied greatly in included studies, daily quantity was divided into the following levels according to the midpoint value of each category: the highest quantity (>3000ml), the 2nd highest quantity (2500-3000ml) and the 3rd highest quantity (<2500ml). The summary OR of these three subgroups were 1.60 (95% CI: 1.17-2.20, P<0.001), 1.24 (95% CI: 0.92-1.67, P=0.16) and 1.07 (95% CI: 0.91-1.27, P=0.42) respectively (Figure 5c). Significant association existed when daily quantity >3000ml.

Dose-response meta-analysis of these 6 studies showed no evidence of nonlinear relationship (p=0.772). Linear analysis suggested that every additional 1000ml consumption would increase the OR by 14.9% (OR=1.149, 95% CI=0.996-1.326) but without statistic significance (p=0.057) (Figure 6b).

Publication bias

There was no asymmetry in funnel plot for the association between total fluid intake and risk of bladder cancer (Supplementary Figure 3). P values for Begg's adjusted rank correlation test was 0.422, and the Egger's regression asymmetry test was 0.243, suggesting that no significant publication bias existed in this meta-analysis.

DISCUSSION

To our knowledge, the current study represents the most comprehensive and up-to-date review on this topic. In this review and meta-analysis of 26 studies with detailed stratified analyses, we showed that there was no unified association between total fluid consumption and the risk of bladder cancer for all people. Nevertheless, high fluid intake would obviously increase the risk in European male and this relationship might also lie in American residents but was relatively weak. On the contrary however, a protective effect from high fluid intake against bladder cancer was noticed in Asia population, especially in Pakistan. The results indicated that the association between total fluid intake and risk of bladder cancer was largely affected by differences in gender, geography region and even the time when the participants were included. Basing on the findings of the current study and those of other previous related studies, we inferred that it was the source and quality of water as well as diets that much related to bladder cancer risk while volume of fluids consumption played a less important role. Our study also provided the evidence that male would be more easily affected by these environmental and living habit factor than female. To prevent higher risk, according to the evidences found in the current study, it is suggested to limit total daily fluid consumption to 2000ml for European males and 3000ml for America residents. There is a significant relationship that every increment 1000ml daily consumption would increase the risk by 28.6% in European male. Similarly every additional 1000ml consumption may increase the OR by 14.9% in American people but the association is not that strong.

What is noteworthy, however, is that the majority of data from Europe and America was collected from cases included before 1990, which may be unable to represent the present conditions. Subgroup analysis by time in the current study indicated that when pooled together, cases recruited before 1990 showed positive outcomes while more recent studies provided negative ones. This difference may mainly generate from the evolution in disinfection techniques and regulations. Taking America for example, since the first regulation on disinfection by-products was passed in 1979, chemical substances including trihalomethanes and free residual chlorine have been successively monitored [33]. Usage of arsenical also keeps declining over the years [8]. All these substances mentioned above are believed to be potential carcinogens for bladder cancer [34-37]. With the improvement of water quality, the increase risk of bladder cancer brought by water pollution is presently better under control. As a result, for the included studies. the conclusions on association between total water intake and risk of bladder cancer altered from positive to negative as time went on.

Another point to be noted is that significant differences were only observed in case-control studies and the outcome varied among regions. Total water intake had a greater impact on residents in European and America than in Asia. Apart from the potential effects of ethnic difference, diet habits, components of daily fluid intake more specifically, might also played important roles. The Chinses are used to drink boiled water, which may further eliminate organic residues in water and reduce the harm taken by them. Furthermore, as one of the most popular beverages in Asia, tea is proved to have protective effect on several cancers, bladder cancer included [38]. Polyphenols in tea was found able to provide protection against bladder cancer by antioxidant activity [39]. These may partly explain why high fluid intake had little impact or even protective effect on risk of bladder cancer in Asian population. On the other hand, tap water is one of the major source of water for people in America and European countries and it has been proved that intake of tap water may induce increasing risk of bladder cancer [6]. In addition, coffee, as one of the major beverages consumed in Europe and America, may also contribute an increased risk of bladder cancer [26, 40, 41]. It was interesting to notice that Wu et al. [40] demonstrated in their study that male coffee drinkers were more likely to develop bladder cancer and Cantor et al. [41] found the

trend in bladder cancer risk from tap water was significant for men only. These outcomes matched our finding that the males were more susceptible to environmental and diet risk factors.

Recently, Buendia et al. [42] observed a slight decrease in urinary adducts formation and significantly decreased urinary mutagenicity with increasing water intake in male smoker, which firstly proved urogenous contact hypothesis. So for the purpose of preventing bladder cancer in European and American population, apart from limiting intake volume as mentioned above without changing drinking style, try to drink more nontap water as well as tea may also help.

Altieri et al. [43] and Bai et al. [7] also conducted meta-analysis on the same topic previously. Altieri et al. explored fluid intake and risk of other cancers as well which made the topic more extensive. But since there were at least 11 more related studies published after his work, his conclusion may unable to keep up to date. Bai et al. conducted their study in 2014 and their did much work on subgroup meta-analysis for different types of beverages and smoke status. However, they missed some important published data and just concluded a sweeping statement but didn't give any specific advice. Our study identified significantly 5 more studies, including those missed by Bai and newly published ones. Because the updated studies provided no detailed data on specific beverages or smoke status, we didn't repeated Bai's work. We focused on subgroup analysis by different exposure level in different gender of population in different region and finally summarized a recommendation list of total fluid intake for preventing increasing risk of bladder cancer.

The limitations of this meta-analysis should be acknowledged. First, only published data were included, making publication bias particularly inevitable even though no significant evidence was observed by statistic analysis. Second, we found that the majority of the data in included studies were collected from cohort or diagnosed cases recruited over 25 years ago, which may unable to reflect the exact conditions of the present. More welldesigned studies are warranted to draw more proper conclusions.

In conclusion, the current study indicates that high fluid intake seems to be a risk factor for bladder cancer in European male and American residents but a protective effect in Asian. Geography region seem to be stronger influence factor on fluid intake and risk of bladder cancer. The reason may lie in differences in water quality and drinking habits. Male are proved to be more susceptible to these potential risk factors. For protective purpose, a recommended total fluid consumption is listed as followed: limitation to 3000ml per day for people living in America and <2000ml per day for European males are recommended. Drinking more boiled water and tea may also help. The actual recommended volume may be higher than mentioned above because a large part of the included data were collected in the past time when water quality wasn't managed as well as the present. So more new casecontrol and cohort studies are needed to draw more proper conclusion.

MATERIALS AND METHODS

The study was designed under the guidance of the Meta-Analysis of Observational Studies in Epidemiology guidelines for study reporting [44].

Literature search

We searched the Web of Science, PubMed and EMBASE to identify articles published up to June 2016. Search terms included "fluid or water" and "intake or consumption" and "bladder cancer or urothelial cancer or transitional or bladder neoplasm or bladder carcinoma". No language limitations were imposed.

Inclusion criteria

Inclusion criteria were as followed: (1) case-control or cohort study exploring the relationship between total fluid consumption and bladder cancer risk, (2) provided exact data in both case and control groups (participants for cohort studies) and (3) adjusted effect estimates with their 95% CIs were directly given or could be calculated from the present data.

Studies with overlapping or insufficient data were excluded.

Selection and data extraction

Two authors (LQ and TY) performed the selection work. Screening was done by titles and abstracts reading. Full text were reviewed when the abstract was insufficient to determine if the study met the inclusion or exclusion criteria. The references of all studies included were manually searched to identify additional studies. Final agreement on inclusion was made by consensus with all authors. All data, including evaluation of study characteristics, risk of bias, and outcome measures, were extracted by two authors.

Methodological quality assessment

The Newcastle-Ottawa Scale (NOS) [45] was used to assess the quality of included studies. Two authors scored these studies independently and final result was made by discussion and consensus with a third author if disagreement came up. Studies that scored >7 were considered as having low risk of bias while scores of <5 indicated high risk of bias.

Statistical analysis

The ORs were extracted and transformed to their natural logs for each include study. The log ORs were weighted by the inverse of their variances to obtain a pooled OR with 95% CI. The Q-test was adopted for statistical heterogeneity measured [46] and the I² score calculating [47]. Heterogeneity was considered present when P<0.10. I² >50% was considered as presence of heterogeneity. In cases lacking of heterogeneity, the Mantel-Haenszel fixed-effect model was used to provide summary estimations, otherwise, the DerSimonian and Laird random-effect model was adopted for the meta-analysis [46, 48]. A dose-response meta-analysis was done according to given data.

Publication bias was assessed through funnel plots as well as tests of Begg rank correlation and Egger regression asymmetry. P<0.05 was considered to be representative of a significant statistical publication bias. Subgroup analyses were conducted as followed to explore the potential heterogeneity among studies: gender (male and female), study design (cohort and case-control studies), geography region (Europe, America and Asia), cases recruited year (before 1990 and after 1990), Fluid intake quantity (selected based on the characteristics of included studies).

Statistical significance was set at P<0.05. STATA version 12.0 (StataCorp, College Station, Texas, United States) and Review Manager 5.2 (The Nordic Cochrane Centre, The Cochrane Collaboration, 2012) was used for the statistical analyses.

CONFLICTS OF INTEREST

All authors declared no conflicts of interest.

GRANT SUPPORT

The current study was supported by the following funds: Technology Support Plan of Science and Technology Department of Sichuan Province (Grant No. 2014SZ0210), Foundation of Sichuan University for Outstanding Youth (Grant No. 2014SCU04B21), Foundation for Academic Leader Fostering of Personnel Department of Sichuan Province (Grant No. JH2014053), Key Project for Applied Research of Organization Department of Sichuan Provincial Party Committee (Grant No. JH2015017) and Natural Science Foundation of China (Grant No. 81470927).

REFERENCES

1. McDonald DF, Lund RR. The role of the urine in vesical neoplasm. I. Experimental confirmation of the urogenous theory of pathogenesis. J Urol. 1954; 71:560-70.

- 2. Oyasu R, Hopp ML. The etiology of cancer of the bladder. Surg Gynecol Obstet. 1974; 138:97-108.
- Silverman DT, Hartge P, Morrison AS, Devesa SS. Epidemiology of bladder cancer. Hematol Oncol Clin North Am. 1992; 6:1-30.
- Michaud DS, Spiegelman D, Clinton SK, Rimm EB, Curhan GC, Willett WC, Giovannucci EL. Fluid intake and the risk of bladder cancer in men. N Engl J Med. 1999; 340:1390-7.
- Burger M, Catto JW, Dalbagni G, Grossman HB, Herr H, Karakiewicz P, Kassouf W, Kiemeney LA, La Vecchia C, Shariat S, Lotan Y. Epidemiology and risk factors of urothelial bladder cancer. Eur Urol. 2013; 63:234-41.
- Villanueva CM, Cantor KP, King WD, Jaakkola JJ, Cordier S, Lynch CF, Porru S, Kogevinas M. Total and specific fluid consumption as determinants of bladder cancer risk. Int J Cancer. 2006; 118:2040-7.
- Bai Y, Yuan H, Li J, Tang Y, Pu C, Han P. Relationship between bladder cancer and total fluid intake: a metaanalysis of epidemiological evidence. World J Surg Oncol. 2014; 12:223.
- Baris D, Waddell R, Beane Freeman LE, Schwenn M, Colt JS, Ayotte JD, Ward MH, Nuckols J, Schned A, Jackson B, Clerkin C, Rothman N, et al. Elevated bladder cancer in Northern New England: the role of drinking water and arsenic. J Natl Cancer Inst. 2016; 108.
- 9. Zhou J, Kelsey KT, Giovannucci E, Michaud DS. Fluid intake and risk of bladder cancer in the Nurses' Health Studies. Int J Cancer. 2014; 135:1229-37.
- Wang J, Wu X, Kamat A, Barton Grossman H, Dinney CP, Lin J. Fluid intake, genetic variants of UDPglucuronosyltransferases, and bladder cancer risk. Br J Cancer. 2013; 108:2372-80.
- Zhou J, Smith S, Giovannucci E, Michaud DS. Reexamination of total fluid intake and bladder cancer in the Health Professionals Follow-up Study Cohort. Am J Epidemiol. 2012; 175:696-705.
- 12. Ahmad MR, Pervaiz MK, Pervaiz G. Non-occupational risk factors of urinary bladder cancer in Faisalabad and Lahore, Pakistan. J Pak Med Assoc. 2012; 62:236-9.
- Ahmad MR, Pervaiz MK, Chawala JA. Risk factors for urinary bladder cancer in Baluchistan. J Ayub Med Coll Abbottabad. 2012; 24:135-7.
- Zhang W, Xiang YB, Fang RR, Cheng JR, Yuan JM, Gao YT. [Total fluid intake, urination frequency and risk of bladder cancer: a population-based case-control study in urban Shanghai]. Zhonghua Liu Xing Bing Xue Za Zhi. 2010; 31:1120-4.
- 15. Hemelt M, Hu Z, Zhong Z, Xie LP, Wong YC, Tam PC, Cheng KK, Ye Z, Bi X, Lu Q, Mao Y, Zhong WD, et al. Fluid intake and the risk of bladder cancer: results from the South and East China case-control study on bladder cancer. Int J Cancer. 2010; 127:638-45.

- Ahmad MR, Pervaiz MK. Risk factors of urinary bladder cancer in Peshawar region of Khyber Pukhtoonkhawa. J Ayub Med Coll Abbottabad. 2010; 22:160-3.
- Jiang X, Castelao JE, Groshen S, Cortessis VK, Shibata DK, Conti DV, Gago-Dominguez M. Water intake and bladder cancer risk in Los Angeles County. Int J Cancer. 2008; 123:1649-56.
- Michaud DS, Kogevinas M, Cantor KP, Villanueva CM, Garcia-Closas M, Rothman N, Malats N, Real FX, Serra C, Garcia-Closas R, Tardon A, Carrato A, et al. Total fluid and water consumption and the joint effect of exposure to disinfection by-products on risk of bladder cancer. Environ Health Perspect. 2007; 115:1569-72.
- Baena AV, Allam MF, Del Castillo AS, Diaz-Molina C, Requena Tapia MJ, Abdel-Rahman AG, Navajas RF. Urinary bladder cancer risk factors in men: a Spanish casecontrol study. Eur J Cancer Prev. 2006; 15:498-503.
- Radosavljevic V, Jankovic S, Marinkovic J, Djokic M. Fluid intake and bladder cancer. A case control study. Neoplasma. 2003; 50:234-8.
- 21. Zeegers MP, Dorant E, Goldbohm RA, van den Brandt PA. Are coffee, tea, and total fluid consumption associated with bladder cancer risk? Results from the Netherlands Cohort Study. Cancer Causes Control. 2001; 12:231-8.
- 22. Geoffroy-Perez B, Cordier S. Fluid consumption and the risk of bladder cancer: results of a multicenter case-control study. Int J Cancer. 2001; 93:880-7.
- 23. Pohlabeln H, Jockel KH, Bolm-Audorff U. Nonoccupational risk factors for cancer of the lower urinary tract in Germany. Eur J Epidemiol. 1999; 15:411-9.
- 24. Bruemmer B, White E, Vaughan TL, Cheney CL. Fluid intake and the incidence of bladder cancer among middleaged men and women in a three-county area of western Washington. Nutr Cancer. 1997; 29:163-8.
- Wilkens LR, Kadir MM, Kolonel LN, Nomura AM, Hankin JH. Risk factors for lower urinary tract cancer: the role of total fluid consumption, nitrites and nitrosamines, and selected foods. Cancer Epidemiol Biomarkers Prev. 1996; 5: 161-6.
- Vena JE, Graham S, Freudenheim J, Marshall J, Zielezny M, Swanson M, Sufrin G. Drinking water, fluid intake, and bladder cancer in western New York. Arch Environ Health. 1993; 48:191-8.
- 27. Kunze E, Chang-Claude J, Frentzel-Beyme R. Life style and occupational risk factors for bladder cancer in Germany. A case-control study. Cancer. 1992; 69:1776-90.
- 28. Slattery ML, West DW, Robison LM. Fluid intake and bladder cancer in Utah. Int J Cancer. 1988; 42:17-22.
- 29. Jensen OM, Wahrendorf J, Knudsen JB, Sørensen BL. The Copenhagen case–control study of bladder cancer. II. Effect of coffee and other beverages. Int J Cancer. 1986; 37:651-7.
- 30. Claude J, Kunze E, Frentzel-Beyme R, Paczkowski K, Schneider J, Schubert H. Life-style and occupational risk

factors in cancer of the lower urinary tract. Am J Epidemiol. 1986; 124:578-89.

- 31. Bianchi GD, Cerhan JR, Parker AS, Putnam SD, See WA, Lynch CF, Cantor KP. Tea consumption and risk of bladder and kidney cancers in a population-based case–control study. Am J Epidemiol. 2000; 151:377-83.
- 32. Ros MM, Bas Bueno-de-Mesquita HB, Buchner FL, Aben KK, Kampman E, Egevad L, Overvad K, Tjonneland A, Roswall N, Clavel-Chapelon F, Kaaks R, Chang-Claude J, et al. Fluid intake and the risk of urothelial cell carcinomas in the European Prospective Investigation into Cancer and Nutrition (EPIC). Int J Cancer. 2011; 128:2695-708.
- Islam N, Sadiq R, Rodriguez MJ, Legay C. Assessing regulatory violations of disinfection by-products in water distribution networks using a non-compliance potential index. Environ Monit Assess. 2016; 188:304.
- 34. Costet N, Villanueva CM, Jaakkola JJ, Kogevinas M, Cantor KP, King WD, Lynch CF, Nieuwenhuijsen MJ, Cordier S. Water disinfection by-products and bladder cancer: is there a European specificity? A pooled and metaanalysis of European case-control studies. Occup Environ Med. 2011; 68:379-85.
- 35. Christoforidou EP, Riza, Kales SN, Hadjistavrou K, Stoltidi M, Kastania AN, Linos A. Bladder cancer and arsenic through drinking water: a systematic review of epidemiologic evidence. J Environ Sci Health A Tox Hazard Subst Environ Eng. 2013; 48:1764-75.
- 36. Hrudey SE, Backer LC, Humpage AR, Krasner SW, Michaud DS, Moore LE, Singer PC, Stanford BD. Evaluating evidence for association of human bladder cancer with drinking-water chlorination disinfection by-products. J Toxicol Env Heal B Crit Rev. 2015; 18:213-41.
- Castaño-Vinyals G, Cantor KP, Grimalt JO, Malats N, Silverman D, Tardon A, Garcia-Closas R. Bladder cancer and exposure to water disinfection by-products through ingestion, bathing, showering, and swimming in pools. Am J Epidemiol. 2007; 165:148-56.
- Zhang YF, Xu Q, Lu J, Wang P, Zhang HW, Zhou L, Ma XQ, Zhou YH. Tea consumption and the incidence of cancer: a systematic review and meta-analysis of prospective observational studies. Eur J Cancer Prev. 2015; 24:353-62.
- 39. Yang CS, Wang X, Lu G, Picinich SC. Cancer prevention by tea: animal studies, molecular mechanisms and human relevance. Nat Rev Cancer. 2009; 9:429-39.
- 40. Wu W, Tong Y, Zhao Q, Yu G, Wei X, Lu Q. Coffee consumption and bladder cancer: a meta-analysis of observational studies. Sci Rep. 2015; 5:9051.
- Cantor KP, Lynch CF, Hildesheim ME, Dosemeci M, Lubin J, Alavanja M, Craun G. Drinking water source and chlorination byproducts. I. Risk of bladder cancer. Epidemiology. 1998; 9:21-8.
- 42. Buendia Jimenez I, Richardot P, Picard P, Lepicard EM, De Meo M, Talaska G. Effect of increased water intake on

urinary DNA adduct levels and mutagenicity in smokers: a randomized study. Dis Markers. 2015; 2015:478150.

- 43. Altieri A, La Vecchia C, Negri E. Fluid intake and risk of bladder and other cancers. Eur J Clin Nutr. 2003; 57:S59-68.
- 44. Stroup DF, Berlin JA, Morton SC, Olkin I, Williamson GD, Rennie D, Moher D, Becker BJ, Sipe TA, Thacker SB. Meta-analysis of observational studies in epidemiology: a proposal for reporting. Meta-analysis Of Observational Studies in Epidemiology (MOOSE) group. JAMA. 2000; 283:2008-12.
- 45. Wells GA, Shea B, O'Connell D, Peterson J, Welch V, Losos M, Tugwell P. The Newcastle-Ottawa Scale (NOS)

for assessing the quality of nonrandomized studies in meta-analysis. [http://www.ohri.ca/programs/clinical_epidemiology/oxford.asp].

- DerSimonian R, Laird N. Meta-analysis in clinical trials. Control Clin Trials. 1986; 7:177-88.
- Higgins JP, Thompson SG, Deeks JJ, Altman DG. Measuring inconsistency in meta-analyses. BMJ. 2003; 327:557-60.
- Mantel N, Haenszel W. Statistical aspects of the analysis of data from retrospective studies of disease. J Natl Cancer Inst. 1959; 22:719-48.