Research Paper

The platelet membrane glycoprotein VI genetic polymorphism (rs1613662, 13254T>C) is not associated with the risk of coronary artery disease

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ABSTRACT

The platelet membrane glycoprotein VI (GP VI), encoded by GP6 gene, is the essential platelet collagen receptor and medicates platelet activation, adhesion and aggregation. Numerous studies revealed that the GP6 genetic polymorphisms may be associated with the susceptibility of coronary artery disease (CAD). However, a clear consensus has not yet been established. To investigate the association between GP6 genetic polymorphisms and CAD, the databases Pubmed, Embase, Chinese National Knowledge Infrastructure (CNKI), and Wanfang were searched for related studies. The pooled odds ratios (ORs) and their corresponding 95% confidence intervals (CIs) were used to evaluate the strength of the association by using a random or fixed-effect model. Our analysis confirmed that there was no significant association between the GP6 13254T>C (Ser219Pro, rs1613662) genetic polymorphism and the risk of CAD under an allelic genetic model (OR = 1.00, 95% CI = 0.79-1.27; P = 0.988), a homozygous genetic model (OR = 1.14, 95% CI = 0.73-1.80; P = 0.563), a heterozygous genetic model (OR = 1.12, 95% CI = 0.95-1.33; P = 0.183), a recessive genetic model (OR = 1.11, 95% CI = 0.71-1.74; P = 0.652). Sensitivity and subgroup analysis indicated the robustness of the results. No publication bias existed between studies. In conclusion, no significant associations between GP6 13254T>C genetic polymorphism and CAD risk were found in this meta-analysis. More large-scale studies on the association of other GP6 genetic polymorphisms and the risk of CAD are needed to be performed in the future.

INTRODUCTION

Cardiovascular diseases are the leading burden of morbidity and mortality worldwide, and coronary artery disease (CAD) accounts for the greatest proportion of cardiovascular diseases [1]. The major forms of CAD include: stable angina, unstable angina, acute coronary syndromes (ACS), myocardial infarction (MI). CAD has a complex pathophysiology determined by lifestyle, environmental and genetic factors [2]. The lifestyle- and environmental–related factors include cigarette smoking, salt intake, obesity, diet, hypertension, diabetes mellitus and other factors [3]. In the past decades, the genetic mechanisms underlying CAD predisposition are widely investigated by case-control association studies [4–11].

Patelet aggregation and thrombosis are involved in the pathogenesis of CAD. In platelets, membrane glycoproteins receptors play crucial roles in adhesion, activation and aggregation, a sequence of events resulting in thrombus formation. Glycoprotein VI (GPVI, GP6), a critical platelet membrane glycoprotein of the immunoglobulin (Ig) superfamily, is an essential receptor for collagen and medicates collagen-induced platelet aggregation and thrombus formation [12, 13]. It is a 60 to 65 kDa type I transmembrane platelet glycoprotein and contains two extracellular Ig-like domains. It forms a complex with the Fc receptor gamma-chain [14]. Upon blood vessel injury, the complex initiates the platelet activation signaling cascade through the exposed subendothelial collagen binding. The platelet-collagen interactions are associated with cardio- and cerebral-vascular diseases in pathologic conditions, and compounds targeting the GPVI-collagen axis have antiatherothrombotic potential [15, 16].

The GP6 gene is mapped on the chromosome 19q13.42 and contains 8 exons spanning over 23 kbp [17]. In consideration of its vital roles in the platelet activation and aggregation, the GP6 gene is considered as an excellent candidate gene for association study in the CAD patients. A plethora of case-control studies have been conducted in order to identify the association between the GP6 13254T>C (Ser219Pro, rs1613662) gene polymorphism and CAD, but contradictory results have been published. Some studies have been reported that the GP6 13254C polymorphism was significantly associated with the susceptibility to CAD, and the GP6 13254C allele increased the risk of CAD [18]. On the contrary, Qin et al. and Yu et al. failed to observe a significant association between the GP6 13254C variant and CAD risk [19, 20]. He et al. also found a similar result in another population [21].

To address the current discordance in the previous findings, we sought to conduct a comprehensive metaanalysis to improve the estimation of association between the GP6 genetic polymorphisms and CAD in the population studies.

RESULTS

Characteristics of eligible studies

Our current meta-analysis was performed according to guidelines of the Meta-analysis of Observational Studies in Epidemiology (MOOSE) statement [22]. A total of 317 relevant papers were produced by our initial literature search, among which 11 studies met the inclusion criteria, including 2692 cases and 2838 controls. As is depicted in the flow diagram (Figure 1), 57 articles were removed because of duplicates and then 201 papers were excluded owing to the obvious irrelevance. The full texts of the remaining 59 articles were reviewed and 10 papers were eligible [18, 19, 21, 23–29]. The paper of Croft SA et al. [18] contained two studies conducted in Sheffield and Leicester, and the data in the paper were extracted seperately for each studies. Therefore, a total of 11 studies for the association between GP6 13254T>C polymorphism and CAD risk were included in our metaanalysis. The information collected from the selected studies was presented in the Table 1. Those countries were included in the meta-analysis as following: United Kingdom, Finland, Czech, Netherlands, USA, Iran and China.

Results of meta-analysis

There was no significant association between the *GP6* 13254T>C genetic polymorphism and CAD in the total population under an allelic genetic model (OR = 1.00, 95% CI = 0.79–1.27; P = 0.988, Table 2 and Figure 2), a homozygous genetic model (OR = 1.14, 95% CI = 0.73–1.80; P = 0.563), a heterozygous genetic model (OR = 1.12, 95% CI = 0.95–1.33; P = 0.183), a recessive genetic model (OR = 1.11, 95% CI = 0.71–1.74; P = 0.652). However, significant association between *GP6* 13254T>C genetic polymorphism and CAD was found under a dominant genetic model (OR = 1.17, 95% CI = 1.01–1.37; P = 0.041, Table 2 and Figure 3)

In the subgroup analysis, there was no significant association in the Chinese population under the allelic, heterozygous and dominant genetic models (P > 0.05, Table 2). Owing to no individuals carrying the CC genotype, meta-analysis could not be made in the Chinese population under the recessive and homozygous genetic models. Additionally, no significant association was also observed in the non-Chinese population under all of the genetic models (P > 0.05) including the dominant genetic model (OR = 1.17, 95% CI = 0.997–1.37; P = 0.054, Table 2 and Figure 3).

Sources of heterogeneity

No significant heterogeneity in the overall population was found under all of the genetic models ($P_{\text{heterogeneity}} > 0.05$, $I^2 < 50\%$) except under the allelic genetic model ($P_{\text{heterogeneity}} = 0.01$, $I^2 = 59.9\%$). Subgroup analysis also showed that heterogeneity was only detected in the non-Chinese subgroup (minor allele frequency, MAF > 0.05) under the allelic genetic model ($P_{\text{heterogeneity}} = 0.002$, $I^2 = 76.8\%$). Therefore, ethnicity was the main confounding factor that could explain the heterogeneity between studies.

Sensitivity analysis

The influence of individual study on the pooled ORs and 95% CIs was evaluated by excluding one single study each time. The corresponding combined ORs and 95% CIs were not significantly altered in the allelic (Figure 4A), recessive (Figure 4B) and homozygous (Figure 4C) genetic models, suggesting a high stability of our meta-analysis results. However, under a dominant model, the significant association (OR = 1.17, 95% CI = 1.01-1.37; P = 0.041) was lost if the study by Croft(Sheffield) *et al.* [18], Ollikainen *et al.* [23], Motovska *et al.* [25], Kazemi *et al.* [28], He *et al.* [21], was removed (Table 3, Figure 4D). In addition, the omission of study by Croft(Leicester) *et al.*

[18] led to a significant association under a heterozygous genetic model (Table 3, Figure 4E).

Publication bias

We performed the Egger's test and Begg's funnel plot to assess the publication bias of the included articles. The *p*-values for Egger's test and Begg's test were summarized in Table 4, which demonstrated that there was no publication bias of the current meta-analysis in all genetic models. And the shape of funnel plot by using allelic genetic model was consistent with these results (Figure 5).

DISCUSSION

The pivotal role of GP6 in platelet activation and aggregation raises the question as to whether the *GP6*

genetic polymorphisms contribute to the susceptibility of CAD. Several GP6 genetic polymorphisms were investigated in the pathogenesis of CAD. To enhance the statistical power, we conducted the present meta-analysis by using the previously published data. The reported genetic polymorphisms include T13254C, A19871G, A21908G, A22630T, C22644A and so on [18]. After the GP6 13254T>C polymorphism was firstly reported to confer an increased risk of MI in the UK, Takagi et.al found that another GP6 genetic polymorphism (C645213T) could also affect the occurrence of myocardial infarction in a Japanese population [30]. With respect to the GP6 13254T>C variant, it is guite paradoxical that the results range from association as a protective factor, to no association, to association as a positive risk factor. Owing to the limited studies on other GP6 genetic polymorphisms, meta-analysis was only conducted in the studies of GP6 13254T>C polymorphism.

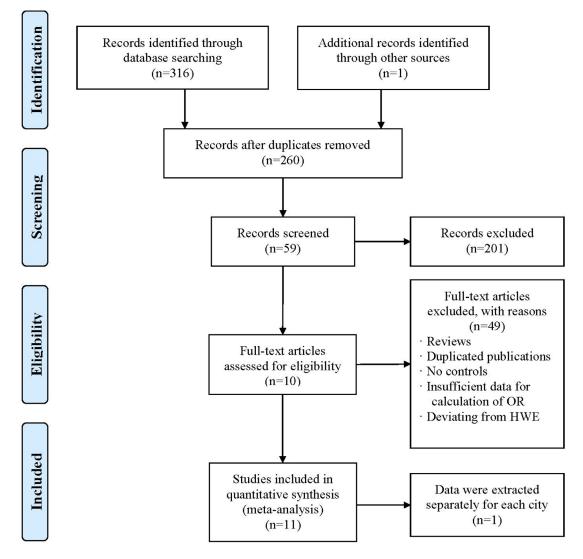


Figure 1: Flow diagram of the search strategy and study selection The terms "*n*" in the boxes represent the number of corresponding studies.

Table 1: Characteristics of the included studies of the association between the GP6 13254T>C genetic polymorphism and coronary artery disease

First	V	Desien	Age (ye	ears old)	Genotyping	Sample size			Case				(Contro	l		MAF	Diagnostic criteria
author	Year	Region	Case	Control	method (c	(case/control)	TT	СТ	CC	Т	С	TT	СТ	CC	Т	С	(control)	(defnition)
Croft	2001	UK (Sheffield)	61.9 ± 9.2	61.1 ± 9.1	PCR-RFLP	289/292	189	89	11	467	111	213	76	3	502	82	0.140	WHO(MI)
Croft	2001	UK (Leicester)	61.5 ± 9.3	54.4 ± 11.8	PCR-RFLP	236/182	166	60	10	392	80	118	57	7	293	71	0.195	WHO(MI)
Ollikainen	2004	Finland	NA	NA	PCR-RFLP	67/250	50	NA	NA	NA	NA	200	48	2	448	52	0.104	NA(MI)
Kou	2004	China	60.0 ± 6.3	58.8 ± 7.3	PCR-RFLP	121/154	112	9	0	233	9	145	9	0	299	9	0.029	WHO(CAD)
Qin	2005	China	59.8 ± 8.9	57.2 ± 9.6	PCR-RFLP	179/164	170	9	0	349	9	155	9	0	319	9	0.027	WHO(ACS)
Motovska	2010	Czech	47.8 ± 6.1	63.66 ± 9.47	PCR-RFLP	105/137	72	30	3	174	36	105	NA	NA	NA	NA	NA	ACC/ESC(MI)
Snoep	2010	Netherlands	57 (49-64)	59 (50-66)	Taqman	547/646	366	163	16	895	195	446	176	21	1068	218	0.170	NA(MI)
Shaffer	2011	USA	NA	NA	Taqman	652/625	NA	NA	NA	1131	173	NA	NA	NA	1024	226	0.181	WHO(MI)
Kazemi	2012	Iran	46.3 ± 5.2	44.7 ± 6.8	PCR-RFLP	100/100	62	34	4	158	42	67	26	7	160	40	0.200	ACC/ESC(MI)
Sun	2012	China	76.0 ± 7.5	71.72 ± 8.3	Sequenom	246/185	240	6	0	486	6	180	5	0	365	5	0.014	NA(CAD)
He	2014	China	66 (45-78)	64 (46-81)	ABI3730XL	150/153	146	4	0	296	4	153	0	0	306	0	0.000	WHO(CAD)

ACC, American Outge of Cardiology; ESC, European Society of Cardiology; ACS, acute coronary synarome; CAJ, contary anter usease; MAF, minor anter requency; Mi, myocardian infarci available; PCR-RELP, polymerase chain reaction-restriction fragment length polymorphism; UK, United Kingdom; USA, United States of America; WHO, World Health Organization.

Table 2: Summary of meta-analysis of association of GP6 13254T>C genetic polymorphism and coronary artery disease

Genetic model	Pooled OR (95% CI)	Z-value	<i>P</i> -value	Literature number	Model	$\mathbf{P}_{\text{heterogeneity}}$	<i>I</i> ² %
Allelic genetic model	1.00 (0.79–1.27)	0.02	0.988	9	R	0.010	59.90%
Chinese subgroup (MAF < 0.05)	1.21 (0.70-2.11)	0.68	0.495	4	F	0.482	0.00%
Non-Chinese subgroup (MAF > 0.05)	0.98 (0.75-1.29)	0.13	0.897	5	R	0.002	76.80%
Recessive genetic model	1.11 (0.71–1.74)	0.45	0.652	4	F	0.164	41.20%
Chinese subgroup (MAF < 0.05)	NA	NA	NA	0	NA	NA	NA
Non-Chinese subgroup (MAF > 0.05)	1.11 (0.71–1.74)	0.45	0.652	4	F	0.164	41.20%
Dominant genetic model	1.17 (1.01–1.37)	2.04	0.041	10	F	0.366	8.30%
Chinese subgroup (MAF < 0.05)	1.22 (0.70-2.13)	0.69	0.49	4	F	0.477	0.00%
Non-Chinese subgroup (MAF > 0.05)	1.17 (0.997–1.37)	1.93	0.054	6	F	0.194	32.20%
Homozygous genetic model	1.14 (0.73–1.80)	0.58	0.563	4	F	0.163	41.50%
Chinese subgroup (MAF < 0.05)	NA	NA	NA	0	NA	NA	NA
Non-Chinese subgroup (MAF > 0.05)	1.14 (0.73–1.80)	0.58	0.563	4	F	0.163	41.50%
Heterozygous genetic model	1.12 (0.95–1.33)	1.33	0.183	8	F	0.418	1.50%
Chinese subgroup (MAF < 0.05)	1.22 (0.70-2.13)	0.69	0.490	4	F	0.477	0.00%
Non-Chinese subgroup (MAF > 0.05)	1.11 (0.93–1.33)	1.18	0.239	4	F	0.198	35.70%

CI, confidence interval; F, fixed-effects model; MAF, minor allele frequency; NA, not available; OR, odds ratio; R, random-effects model.

In our meta-analysis of the *GP6* 13254T>C genetic polymorphism, there was no significant association between and CAD in the total population under an allelic genetic model, a homozygous genetic model, a heterozygous genetic model, a recessive genetic model. The significant association between *GP6* 13254T>C genetic polymorphism and CAD was only found under a dominant genetic model. However, the significant association was not exited in the subgroups stratified by ethnicity. Therefore, the *GP6* 13254T>C polymorphism was not associated with CAD.

The mature GP6 protein consists of 319 amino acids and the *GP6* gene 13254 T>C (rs1613662) polymorphism is an amino acid substitution of serine 219 by proline (Ser219Pro) in the exon 5, which is a likely factor for the functional differences of two common *GP6* haplotypes (GP6a and GP6b). The amino acid substitution was predicted to affect at least the local secondary/tertiary structure of the GPVI receptor. Previous studies indicated that the GP6 13254 T>C polymorphism played important roles in the expression of receptor, platelet activation and aggregation, signaling and fibrinogen binding [31, 32]. Trifiro et al. found that the GP6 13254 T>C polymorphism or the GP6a/GP6b haplotypes do not influence the GPVI ligand-binding affinity and expression levels [33]. This may partly explain why the association between the GP6 13254 T>C polymorphism and CAD risk was negative in our meta-analysis results. In addition, other genetic variants in the GP6 gene may also affect the expression of GP6 and the platelet function. The genomic structure GP6 is composed of 8 exons and 7 introns. Genetic variations in the regulatory region may alter the sequence of protein and lead to its functional effects. A recent study, conducted by Liu et al. [34], showed that GP6 rs1671153 and rs1654419

genetic polymorphisms were significantly with the risk of CAD. The polymorphisms rs1671153 and rs1654419 were located in the intron 6 and intron 5 respectively, which may influence the splicing of *GP6* and be involved in the regulation of mRNA stability. However, the mechanism needs to be investigated by functional validation of this particular SNP.

Based upon comparison of the MAF value in the included studies, we found there was a racial difference in the distribution of the *GP6* 13254T>C genetic polymorphism. The frequency of the *GP6* 13254C allele in Caucasian American and European population was nearly 20% [18, 26, 27], which is considerably different from the Chinese individuals (less than 5%) [19, 21] and the Japanese population (MAF = 2%) [30]. In addition, The Iranian population (MAF = 20%) [28] is genetically very similar to Caucasian American and European population individuals. Importantly, the difference in allele frequencies among ethnicities is also consistent

with the data from the 1000 genomics database (https:// www.ncbi.nlm.nih.gov/variation/tools/1000genomes/). Therefore, subgroup analysis stratified by the MAF value was conducted in our meta-analysis.

To the best of our knowledge, this is a comprehensive report with the largest sample size to determine the association between *GP6* 13254T>C genetic polymorphism and the susceptibility of CAD by a metaanalysis. The major strength of this current study was lack of significant heterogeneity under all of the genetic models except under the allelic genetic model. The sources of between-study heterogeneity were explored by the subgroup analysis, and the results showed that ethnicity contributed to heterogeneity in the allelic genetic model. Sensitivity analysis also demonstrated that the omission of any individual study did not significantly change the pooled estimates of meta-analysis under the allelic, recessive and homozygous genetic models. Additionally, in the analysis of publication bias, the Begg's funnel plot

Study		%
0	OR (95% CI)	Weight
Ion-Chinese(MAF>0.05)		
Croft(Sheffield) (2001)	1.46 (1.07, 1.99)	17.19
Croft(Leicester) (2001)	0.84 (0.59, 1.20)	15.83
Snoep (2010) 🔶	1.07 (0.86, 1.32)	20.48
Shaffer (2011) 🔶	0.69 (0.56, 0.86)	20.39
Kazemi (2012)	1.06 (0.65, 1.73)	12.04
Subtotal (I-squared = 76.8%, p = 0.002)	0.98 (0.75, 1.29)	85.93
Chinese(MAF<0.05)		
Kou (2004)	1.28 (0.50, 3.28)	5.02
Qin (2005)	0.91 (0.36, 2.33)	5.05
Sun (2012)	0.90 (0.27, 2.98)	3.37
He (2014)	9.30 (0.50, 173.56)	0.63
Subtotal (I-squared = 0.0%, p = 0.482)	1.13 (0.64, 1.99)	14.07
Overall (I-squared = 59.9%, p = 0.010)	1.00 (0.79, 1.27)	100.00
IOTE: Weights are from random effects analysis		
.00576 1	174	

Figure 2: Forest plot for the allelic genetic model stratified by ethnicity (minor allele frequency of *GP6* 13254T>C polymorphism) in studies with coronary artery disease patients.

did not reflect remarkable asymmetry and the results of Egger's test were consistent in all genetic models. Therefore, all these analyses made our meta-analysis findings robust and reliable.

However, our meta-analysis is not without limitations. Firstly, since the common environmental risk factors (e.g. age, gender, obesity, dyslipidemia, hypertension, diabetes or smoking) of the CAD were not available in the included studies, we were unable to obtain these factors to adjust the meta-analysis. Thus, the results in our meta-analysis were based primarily on the crude ORs and its 95% CIs. Secondly, because the included studies were all retrospective case-control association researches, some potential undetected bias may not be excluded. Thirdly, an interaction between hormone replacement therapy and GP6 gene on risk for CAD existed in the previous reports [27], but similar analysis weren't be performed in our article due to lack of the information in most studies. Fourthly, sensitivity analysis showed that the pooled results could be altered by some studies under the dominant and heterozygous

model, and thus these need to be interpreted with caution. Last but not the least, owing to a polygenic disease like CAD, the effect of haplotypes or gene-gene interaction deserved to be investigated. Even though several *GP6* genetic polymorphisms have been associated with CAD [30], the pooled results regarding other *GP6* variants can't be available because of the limited studies.

In conclusion, this meta-analysis showed that the *GP6* 13254T>C (rs1613662) polymorphism was not significantly associated with CAD risk. Given the limitations mentioned above, more large-scale studies on the association of other *GP6* genetic polymorphisms (e.g. rs1671153, rs1654419) and the risk of CAD deserved to be performed to further confirm our findings in the future.

MATERIALS AND METHODS

Search strategy and selection criteria

We performed a systematic published data search for the association studies between *GP6* genetic

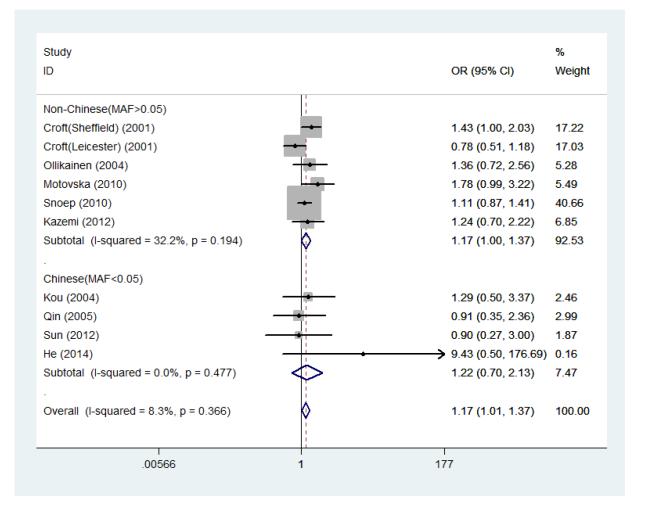


Figure 3: Forest plot for the dominant genetic model stratified by ethnicity (minor allele frequency of *GP6* 13254T>C polymorphism) in studies with coronary artery disease patients.

polymorphisms and CAD, published before May 2017 on the electronic databases PubMed, Embase, Chinese National Knowledge Infrastructure (CNKI), and Wanfang using the following search terms: (coronary artery disease or coronary heart disease atherosclerosis or myocardial infarction or myocardial infarct or heart attack or MI) and (polymorphism or single nucleotide polymorphism or SNP or variant or variation) and (glycoprotein VI or GP VI or GP6 or platelet membrane glycoprotein). No language restrictions were used. All eligible studies were retrieved, and cited references were carefully examined for additional potentially relevant studies.

The included articles in the meta-analysis should conform to the criteria as follows: (a) case-control

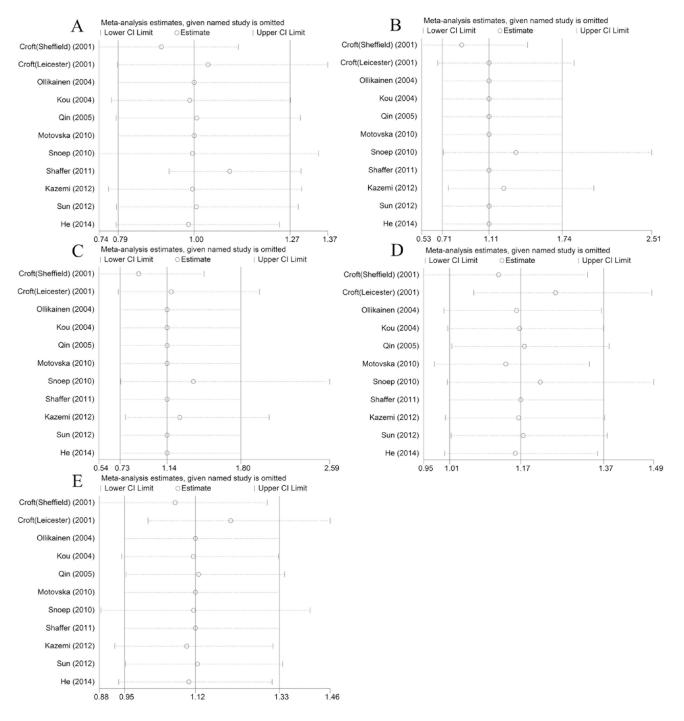


Figure 4: The sensitivity analysis of the pooled ORs and 95% CIs for *GP6* 13254T>C polymorphism under the allelic (**A**) recessive (**B**) homozygous (**C**) dominant (**D**) and heterozygous (**E**) genetic models. CI, confidence interval; OR, odds ratio.

First author	Year	Domin	ant genetic n	nodel	Heterozygous genetic model				
rirst author	rear	Estimate 95% CI		Estimate	95% CI				
Croft (Sheffield)	2001	1.12	0.95	1.33	1.07	0.88	1.30		
Croft (Leicester)	2001	1.26	1.06	1.48	1.21	1.01	1.46		
Ollikainen	2004	1.16	0.99	1.36	1.12	0.95	1.33		
Kou	2004	1.17	1.00	1.37	1.12	0.94	1.33		
Qin	2005	1.18	1.01	1.38	1.13	0.95	1.34		
Motovska	2010	1.14	0.97	1.34	1.12	0.95	1.33		
Snoep	2010	1.22	1.00	1.49	1.12	0.89	1.41		
Shaffer	2011	1.17	1.01	1.37	1.12	0.95	1.33		
Kazemi	2012	1.17	1.00	1.37	1.10	0.92	1.32		
Sun	2012	1.18	1.01	1.38	1.13	0.95	1.34		
Не	2014	1.16	0.99	1.35	1.11	0.93	1.31		
Combined		1.17	1.01	1.37	1.12	0.95	1.33		

Table 3: Sensitivity analysis for the *GP6* 13254T>C genetic polymorphism under the dominant and heterozygous genetic models

95% CI, 95% confidence interval.

design. (b) assessment of the association of GP6 genetic polymorphisms with CAD. c) data on the GP6 genotypes was available in both cases and controls. d) Genotypes in the control subjects should be in agreement with the Hardy-Weinberg equilibrium (HWE). Studies were excluded if any of the following applies: (a) reviews and repeated publications; (b) studies not meeting all of the inclusion criteria. was to resolve the possible discrepancies between the two investigators. The similar data in different studies by the same author group were only adopted once. The following information was drawn out: first author's name, publication year, region, ethnicity, sample size, genotyping method and number of genotype in case and control group.

Statistical analysis

Data extraction

Data were independently extracted according to a standardized protocol by two investigators. The third one

The odds ratio (OR) and their corresponding 95% confidence interval (CI) were used to compare the association between GP6 13254T>C polymorphism and CAD. The pooled ORs were assessed with allelic

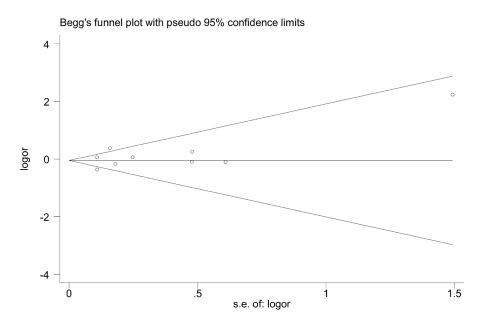


Figure 5: Begg's funnel plot for studies of the association between coronary artery disease and *GP6* 13254T>C polymorphism under an allelic genetic model.

Table 4: Egger's and Begg's test for the publication bias of GP6 13254T>C genetic polymorphism

66 66	1	
Genetic Models	Egger's Test <i>p</i> Value	Begg's Test p Value
Allelic genetic model	0.369	0.754
Dominant genetic model	0.402	0.592
Recessive genetic model	0.614	0.734
Homozygous genetic model	0.596	0.734
Heterozygous genetic model	0.590	0.711

model (C versus T), homozygous model (CC versus TT), heterozygous model (TC versus TT), recessive model (CC versus TC/TT), dominant model (CC/TC versus TT). The Z test was used to determine the pooled ORs with the significance set at P < 0.05. The Chi-square test was used to assess the HWE in the control groups. The betweenstudy heterogeneity was determined by the I^2 statistic test, which is not inherently dependent on the number of studies for the meta-analysis [35]. If obvious heterogeneity existed among the individual studies ($I^2 > 50\%$), the randomeffects model would be used to calculate the pooled OR and its 95% CI [36]. If no heterogeneity is detected, the fixed-effects model using the Mantel-Haenszel method would be adopted for the meta-analysis [37]. Subgroup analysis according to the ethnicity or the MAF value was conducted to evaluate the association and explore the sources of between-study heterogeneity. Sensitivity analysis was performed to assess the effect of individual study on the combined results and evaluate the stability of results. The potential publication bias was detected by the Begg's funnel plot [38], and the funnel plot asymmetry was evaluated with the Egger's linear regression test [39]. The STATA 12.0 software (StataCorp, College Station, TX, USA) was used to perform all statistical analysis.

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CONFLICTS OF INTEREST

The authors declare no competing financial interests.

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