Meta-Analysis

# Association between *TGFB1* genetic polymorphisms and chronic allograft dysfunction: a systematic review and meta-analysis

Kun Liu<sup>1,\*</sup>, Xuzhong Liu<sup>1,\*</sup>, Shuo Gu<sup>1</sup>, Qing Sun<sup>1</sup>, Yunyan Wang<sup>1</sup>, Junsong Meng<sup>1</sup> and Zongyuan Xu<sup>1</sup>

<sup>1</sup>Department of Urology, Huai'an First People's Hospital, Nanjing Medical University, Huai'an, 223300 China

<sup>\*</sup>These authors contributed equally to this work

Correspondence to: Zongyuan Xu, email: w541003954@126.com Junsong Meng, email: w675921038@126.com

Keywords: TGFB1, polymorphism, chronic allograft dysfunction, meta-analysis

Accepted: July 13, 2017

**Received:** June 20, 2017

Published: July 24, 2017

Copyright: Liu et al. This is an open-access article distributed under the terms of the Creative Commons Attribution License 3.0 (CC BY 3.0), which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited.

#### ABSTRACT

Background: Epidemiological studies have investigated the role of transforming growth factor- $\beta$ 1 (TGF- $\beta$ 1) in chronic allograft dysfunction (CAD) following kidney transplantation. *TGFB1* gene polymorphisms (SNP rs1800470 and rs1800471) may be associated with the risk of CAD. In this meta-analysis, the relationship between these two variations and the risk of CAD was explored.

Materials and Methods: MEDLINE, EMBASE, the Cochrane Central Register of Controlled Trials (CENTRAL), Embase, the Chinese CNKI and WANFANG databases were searched. Data were extracted and pooled results were estimated from odds ratios (ORs) with 95% confidential intervals (95% CIs). Quality assessments were performed, and publication bias of all eligible studies examined.

Results: Eight studies with 1038 subjects were included in our analysis. According to the effects on TGF- $\beta$ 1 secretion, haplotypes were categorized as "HIGH", "INTERMEDIATE" and "LOW". The combined results showed a statistically significant difference of *TGFB1* haplotypes between the CAD recipients and control subjects when "HIGH" with "INTERMEDIATE" and "LOW" ("HIGH" vs. "INTERMEDIATE" + "LOW": OR: 3.56, 95% CIs: 2.20, 5.78, P < 0.001) were compared. No significant association was found between the *TGFB1* codon 10 or codon 25 and the CAD risk in all five genetic models.

Conclusions: Our meta-analysis has found the haplotype of *TGFB1* codon 10/25 T/T G/G and T/C G/G genotypes, associated with increased production of TGF- $\beta$ 1, was linked with CAD risk following kidney transplantation. Moreover, no significant difference was found between TGFB1 codon 10 or codon 25 and the development of CAD.

## **INTRODUCTION**

Kidney transplantation is the optimal therapy for end-stage renal disease [1]. Short-term allograft survival has been significantly improved due to advancements in immuosuppressive agents and surgical techniques [2]. However, long-term allograft survival, especially in the period of 10 to 20 years after kidney transplantation, has not accommodated the increase of short-term survival. Numerous studies have shown that chronic allograft dysfunction (CAD) is the main cause of graft failure [3, 4].

CAD is mainly characterized as glomerular sclerosis, tubular atrophy and interstitial fibrosis [5]. Various factors contribute to the pathogenesis of CAD, such as chronic calcineurin inhibitor (CNI), particularly cyclosporine A (CsA) nephrotoxicity, fibrosis-related cytokines and immune-associated factors [6–8]. Among these factors, certain cytokines, including transforming

growth factor- $\beta$ 1 (TGF- $\beta$ 1), tumor necrosis factor- $\alpha$  (TNF- $\alpha$ ) and cell tissue growth factors (CTGF), were considered to be closely associated with the development of fibrogenesis, thus contributing to the pathogenesis of CAD [5, 9, 10]. Recent studies have reported that TGF- $\beta$ 1 could promote the development of interstitial fibrosis and tubular atrophy through epithelial-to-mesenchymal transition (EMT), endothelial-to-mesenchymal transition (EMT) and other mechanisms [5, 11].

TGF-β1 is a multifunctional growth cytokine in humans and one of the significant mediators of wound healing and tissue regeneration [12]. TGFB1 is mapped on to chromosome 19q13.1-13.3 with seven exons and six introns, and the expression of TGFB1 and production of TGF- $\beta$ 1 is related to single nucleotide polymorphisms (SNPs). TGFB1 contains two SNPs, +869T/C at codon 10 (rs1800470) and +915G/C at codon 25 (rs1800471) contributing to variation in TGF-B1 production both in vitro and in serum [13]. There is increasing evidence that the level of production of a series of cytokines can be modulated by polymorphisms in the corresponding genes [14]. Since fibrosis-related cytokines play an important role in the inflammatory and immune response that mediate the outcome of kidney transplantation, the relationship between those cytokine polymorphisms, particularly TGFB1 SNPs, and CAD has been explored in several studies [15-17]. However, the results of these studies were inconsistent and often conflicting.

Based on the crucial role of TGF- $\beta$ 1 in CAD pathogenesis, we performed meta-analysis to investigate the contributions of two TGFB1 SNPs to CAD risk.

# **MATERIALS AND METHODS**

## Search strategy

A comprehensive literature search was performed in PubMed, the Cochrane Central Register of Controlled Trials (CENTRAL), Embase, the Chinese CNKI and WANFANG databases (updated on December 20st 2016) by two independent authors (Kun L and Xuzhong L). The following keywords were used: (transforming growth factor-beta 1 OR TGFB1), AND (polymorphisms OR SNPs OR variants), AND (chronic rejection OR chronic allograft nephropathy OR chronic allograft dysfunction), AND (MESH item, kidney transplantation). The equivalent Chinese terms were used in the Chinese databases. Furthermore, the reference lists of all studies included in the meta-analysis were also reviewed for possible inclusion.

## Inclusion and exclusion criteria

Studies were included if they met the following eligibility criteria: (1) case-control studies designed to investigate the relationship between *TGFB1* SNPs and

chronic allograft dysfunction after kidney transplantation; (2) available information on the frequencies of genotype or allele in case and control groups; (3) all subjects from three allelic groups were derived from a population within the same geographic area and ethnic background; (4) full-text article were published in English or Chinese. The exclusion criteria were: (1) Case reports; (2) reviews; (3) animal experiment, chemistry, or cell line studies; (4) studies in a language other than English or Chinese; (5). No eligible or insufficient data frequencies of genotypes for each polymorphism could be extracted for meta-analysis. Two authors (Kun L and Xuzhong L) independently assessed and selected trials for final analysis with discrepancies resolved by consensus.

## **Data extraction**

Two investigators (Kun L and Xuzhong L) independently extracted relevant data from all selected studies and reached consensus for all items. Basic characteristics of patients were collected as following: first author's name, year of publication, nation/race, number of subjects, the proportion of male subjects, mean age and method of genotype. In addition, the genotype and allele frequencies of *TGFB1* SNPs were collected using a standardized data extraction form. Missing data was examined by contacting the first or corresponding author.

#### Quality assessment

The methodological quality of each included study was assessed by the Newcastle-Ottawa quality assessment scale (NOS)(18). Each study was evaluated on the standard criteria and categorized based on three factors, including selection, comparability and exposure. Scores ranged from 0 to 9; 9 points represented the highest quality and lowest risk of bias.

## Statistical analysis

The pooled data was used to assess the strength of the association between TGFB1 polymorphisms and CAD using odds ratio (OR) with 95% confidence intervals (95% CIs) in a dominant model, a recessive model, a co-dominant model, a co-recessive model and an allele model. A p value less than 0.05 was considered statistically significant. Heterogeneity test among trials was determined by  $I^2$ , and was defined as 100% \*(Q-df )/Q, where Q is Cochran's heterogeneity statistic and df is the degrees of freedom, with a fixed-effect model set at low statistical inconsistency ( $I^2$ < 25%). Otherwise, a random-effects model was selected, which is better adapted to clinical and statistical variations [18–19]. To explore the potential effects of heterogeneity, stratification analysis by ethnicity, age and quality criteria was carried out. The Egger's regression test and funnel plots were used to assess potential publication bias. Cumulative

meta-analysis was carried out by the year of publication. All analyses were performed using STATA (release 12.0, College Station, TX).

## RESULTS

#### Literature selection and study characteristics

A flow diagram of the screening process for included studies is shown in Figure 1. Primary screening identified 18 potentially relevant articles, including 15 articles in English and 3 in Chinese. Based on their titles and abstracts, 8 articles were excluded due to unrelated subject or no available data. After screening the full articles, a total of 8 trials with 1038 subjects met the criteria and were included in our meta-analysis.

Basic characteristics of eligible studies are shown in Table 1. Overall, the meta-analysis of included studies comprised of *TGFB1* codon 10 and codon 25. Four studies [17, 20–22] included patients with CAN, confirmed by renal biopsy. CAD recipients included in another 4 trials were evaluated by the clinical examinations. The 8 studies included 4 Asian [20–23], 3 European [15, 17, 24] and 1 Latino [25] populations. Results of quality assessment using the NOS scale is presented in Table 2.

## Quantitative synthesis

A total of 8 trials were included in the analysis of association between *TGFB1* polymorphisms and CAD. For *TGFB1* codon 10 T/C, no significant difference was found between CAD recipients and control subjects in all

five models (TT vs. TC+CC: OR: 1.37, 95% CIs: 0.61, 3.06, P = 0.44; CC vs. TC+TT: OR: 1.36, 95% CIs: 0.77, 2.40, P = 0.29; TT vs. CC: OR: 1.15, 95% CIs: 0.36, 3.67, P = 0.82; TT vs. TC: OR: 1.24, 95% CIs: 0.52, 2.94, P = 0.63; T vs. C: OR: 1.29, 95% CIs: 0.78, 2.12, P = 0.32. Supplementary Figure 1). For *TGFB1* codon 25 G/C, there was no significant association between CAD recipients and control groups in dominant model, co-dominant model and allele model (GG vs. GC+CC: OR: 1.31, 95% CIs: 0.46, 3.74, P = 0.62; GG vs. GC: OR: 1.36, 95% CIs: 0.49, 3.02, P = 0.68. Supplementary Figure 2).

Furthermore, we performed the meta-analysis to investigate the effect of *TGFB1* haplotype on the pathogenesis of CAD. According to the effects on TGF- $\beta$ 1 secretion, these haplotypes were categorized as "HIGH", "INTERMEDIATE" and "LOW" (Supplementary Table 1) [16]. In summary, we observed statistically significant differences in *TGFB1* haplotypes between the CAD recipients and control subjects when comparing the "HIGH" with "INTERMEDIATE" and "LOW" ("HIGH" vs. "INTERMEDIATE" + "LOW": OR: 3.56, 95% CIs: 2.20, 5.78, P < 0.001, Figure 2).

#### Sensitivity analysis and publication bias

Sensitivity analysis was performed by sequential removal of individual studies to reflect the effect of individual data on the pooled ORs. No effect of any study was found on the pooled results in the above five models. The Begg's funnel plot and Egger's linear regression test was performed to determine publication bias. Results of

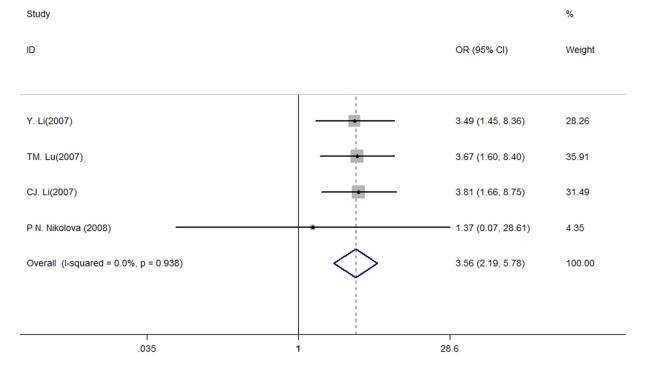


Figure 1: Flow chart of literature search and study selection.

| Author (year)                | Ethnicity | Case number | Age (years; mean ± SD)                    | Male/female | Genotyping method            | Genetic<br>equilibrium | Diagnosis criteria |
|------------------------------|-----------|-------------|---|-------------|------------------------------|------------------------|--------------------|
| A. Melk (2003)               | Caucasian | 105         | Con:8.6 $\pm$ 4.4; Case:9.2 $\pm$ 4.3     | 57/48       | RT-PCR                       | NA                     | CAD                |
| J Y. Park(2004)              | Asian     | 163         | NA  | NA          | PCR-SSCP                     | NA                     | CAN                |
| K. Gendzekhadze(2006)        | Latino    | 62          | $40\pm10$                                 | 37/26       | PCR-SSP                      | Yes                    | CAD                |
| Y. Li(2007)                  | Asian     | 122         | 39.7                                      | 95/27       | PCR                          | NA                     | CAD                |
| TM. Lu(2007)                 | Asian     | 144         | 40.6                                      | 109/35      | PCR-SSP                      | NA                     | CAN                |
| CJ. Li(2007)                 | Asian     | 100         | Con: $42.8 \pm 7.9$ ;Case: $45.3 \pm 6.7$ | 51/49       | PCR                          | NA                     | CAN                |
| P N. Nikolova (2008)         | Caucasian | 66          | NA  | 39/27       | PCR-SSP                      | NA                     | CAN                |
| M A. Jiménez-Sousa<br>(2012) | Caucasian | 276         | Con:50.5 ± 15.2;Case:51.5 ± 18            | 108/168     | the SNPlex genotyping system | Yes                    | CAD                |

Abbreviations: SD, standard deviation; NA, not available; CAD, chronic allograft dysfunction; CAN, chronic allograft nephropathy; PCR-SSCP, PCR-single strand conformation polymorphism; PCR-SSP, PCR-sequence-specific primers.



# **PRISMA 2009 Flow Diagram**

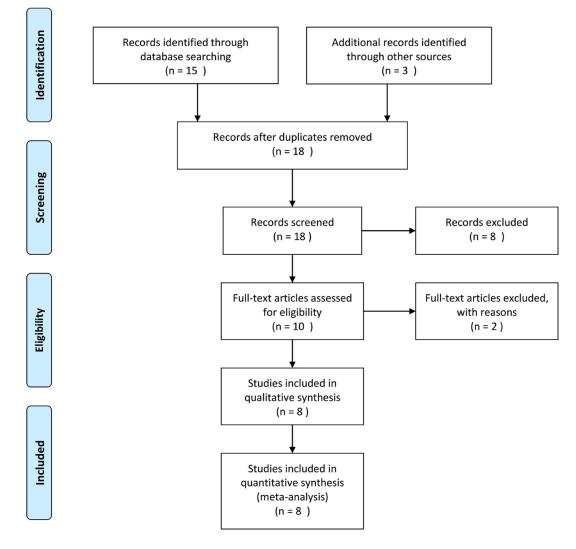


Figure 2: Meta-analysis of the association between the *TGFB1* haplotypes and CAD risk (HIGH vs. INTERMEDIATE + LOW).

| Studies                       | Selection                      |                                    |                          |                           | Comparability         |   | Exposure                     |                                       |                  | Total<br>quality |
|-------------------------------|--------------------------------|------------------------------------|--------------------------|---------------------------|-----------------------|---|------------------------------|---------------------------------------|------------------|------------------|
|                               | Case<br>definition<br>adequate | Representativeness<br>of the cases | Selection of<br>controls | Definition<br>of controls | Adjustment<br>for age | Adjustment<br>for lifestyle/<br>traditional<br>risk factors | Ascertainment<br>of exposure | Uniform<br>method of<br>ascertainment | response<br>rate | score            |
| A. Melk (2003)                | 1                              | 1                                  | 1                        | 1                         | 1                     | 1   | 0                            | 1                                     | 0                | 7                |
| J Y. Park(2004)               | 1                              | 1                                  | 1                        | 1                         | 1                     | 1   | 0                            | 1                                     | 0                | 7                |
| K. Gendzekhadze<br>(2006)     | 1                              | 1                                  | 1                        | 1                         | 1                     | 1   | 0                            | 1                                     | 0                | 7                |
| Y. Li(2007)                   | 1                              | 1                                  | 0                        | 1                         | 0                     | 1   | 0                            | 1                                     | 0                | 5                |
| TM. Lu(2007)                  | 1                              | 1                                  | 0                        | 1                         | 1                     | 1   | 0                            | 1                                     | 0                | 6                |
| CJ. Li(2007)                  | 1                              | 1                                  | 0                        | 1                         | 0                     | 1   | 0                            | 1                                     | 0                | 5                |
| P N. Nikolova<br>(2008)       | 1                              | 1                                  | 1                        | 1                         | 1                     | 1   | 0                            | 1                                     | 0                | 7                |
| M A. Jiménez-<br>Sousa (2012) | 1                              | 1                                  | 1                        | 1                         | 1                     | 1   | 0                            | 1                                     | 0                | 7                |

Table 2: Newcastle-Ottawa quality assessment scale for each included study

Table 3: Begg's funnel plot and Egger's linear regression test for all included studies

| genotypes                 | Egger's | Begg's test |       |      |
|---------------------------|---------|-------------|-------|------|
|                           | t       | Р           | Z     | Р    |
| Codon 10                  |         |             |       |      |
| TT vs. TC+CC              | -1.15   | 0.33        | 0.49  | 0.62 |
| CC vs. TC+TT              | -0.23   | 0.85        | 0.52  | 0.60 |
| TT vs. TC                 | -1.89   | 0.16        | 0.49  | 0.62 |
| TT vs. CC                 | -0.08   | 0.95        | -0.52 | 0.60 |
| T vs. C                   | 0.26    | 0.81        | -0.49 | 0.62 |
| Codon 25                  |         |             |       |      |
| GG vs. GC+CC              | 0.93    | 0.52        | 0.52  | 0.60 |
| GG vs. GC                 | 0.40    | 0.76        | 0.52  | 0.60 |
| G vs. C                   | 1.00    | 0.50        | 0.52  | 0.60 |
| Haplotype                 |         |             |       |      |
| HIGH vs. INTERMEDIATE+LOW | 2.13    | 0.28        | -0.52 | 0.60 |

funnel plot and regression test showed that there was no significant publication bias (Table 3).

# DISCUSSION

The present meta-analysis included 8 studies with 1038 renal transplant recipients to assess the correlation between two *TGFB1* SNPs, codon 10 and 25, and CAD risk. In addition, this meta-analysis presented the first cumulative meta-analysis on this topic. In the cumulative meta-analysis, we found that "HIGH" haplotype, which contains codon 10, 25 T/T G/G and T/C G/G genotypes, was strongly associated with the development of CAD following kidney transplantation. However, no significant association of *TGFB1* SNPs in all five genetic models with CAD risk was observed.

TGF- $\beta 1$  is a multifunctional cytokine that regulates the proliferation and differentiation of many cell types, and has

been identified as an important promoter of fibrogenesis in various cells and tissues [26]. There is increasing evidence that TGF- $\beta$ 1 is strongly associated with the pathogenesis of interstitial fibrosis in kidney through a physical process, including EMT and EndMT [27]. The *TGFB1* SNPs +869T/C at codon 10 and +915G/C at codon 25 are important in the signal sequence and may influence the secretion of TGF- $\beta$ 1, contributing to the occurrence of CAD [13]. In our study, we found that the haplotype of *TGFB1* codon 10, 25 T/T G/G and T/C G/G genotypes, associated with higher production of TGF- $\beta$ 1, are more susceptible to CAD following kidney transplantation when compared with haplotypes of other genotypes. This is consistent with the results of studies conducted by P N. Nikolova [17].

The *TGFB1* codon 10 and codon 25 were associated with inter-individual variation in the level of TGF- $\beta$ 1 production. Dunning et al. (13) found that secretion of the proline form of TBG- $\beta$ 1 was 2.8 times that of the leucine

form of TGF- $\beta$ 1 at codon 10 in transfected Hela cells, therefore the proline mutation at codon 10 could increase the amount of TGF- $\beta$ 1 protein secretion. Similarly, Cambien et al. [28] believed that the substitution of arginine with proline corresponded to a change from a large polar acid to a small apolar acid and concluded that this may affect the export of TGF- $\beta$ 1 protein. In our study, no significant association of *TGFB1* codon 10 or codon 25 with the CAD risk was found in all five models, suggesting that the single mutation in codon 10 and codon 25 were not responsible for the pathogenesis of CAD after kidney transplantation.

Notably, these results should be interpreted with caution. The case number of recipients with CAD included in our analysis is relatively small, which could lead to relatively high heterogeneity. Due to the limited number of studies in our analysis, we categorized CAN, confirmed by allograft renal biopsy, into the diagnosis of CAD, instead of subgroup analysis. Finally, there were at least 10 SNPs reported in the *TGFB1* gene, and except for codon 10 and codon 25, rare studies have focused on other SNPs, which may play a role in the pathogenesis of CAD.

In conclusion, our meta-analysis found the haplotype of *TGFB1* gene codon 10/25 T/T G/G and T/C G/G genotypes, associated with increased production of TGF- $\beta$ 1, was linked with CAD risk following kidney transplantation. Moreover, no significant difference was found between *TGFB1* codon 10 or codon 25 and the development of CAD. Further studies incorporating subjects with difference ethnic backgrounds combined with re-sequencing of the marked region and functional evaluations are warranted.

# **ACKNOWLEDGMENTS AND FUNDING**

We would like to thank the native English speaking scientists of Elixigen Company (Huntington Beach, California) for editing our manuscript.

# **CONFLICTS OF INTEREST**

The authors have no conflicts of interest to declare.

# REFERENCES

- Wang Z, Han Z, Tao J, Lu P, Liu X, Wang J, Wu B, Huang Z, Yin C, Tan R, Gu M. Clinical efficacy and safety of pamidronate therapy on bone mass density in early post-renal transplant period: a meta-analysis of randomized controlled trials. PloS one. 2014; 9:e108106.
- Djamali A, Reese S, Yracheta J, Oberley T, Hullett D, Becker B. Epithelial-to-mesenchymal transition and oxidative stress in chronic allograft nephropathy. American journal of transplantation. 2005; 5:500–9.
- Transplant U. Renal Transplant Audit 1990–1998. London: UK Transplant. 2001.

- Goldberg RJ, Weng FL, Kandula P. Acute and Chronic Allograft Dysfunction in Kidney Transplant Recipients. The Medical clinics of North America. 2016; 100:487–503.
- Wang Z, Han Z, Tao J, Wang J, Liu X, Zhou W, Xu Z, Zhao C, Ju X, Wang Z, Tan R, Gu M. Transforming Growth Factor-beta1 Induces Endothelial-to-Mesenchymal Transition via Akt Signaling Pathway in Renal Transplant Recipients with Chronic Allograft Dysfunction. Annals of transplantation. 2016; 21:775–83.
- Zhang J, Wang Z, Xu Z, Han Z, Tao J, Lu P, Huang Z, Zhou W, Zhao C, Tan R, Gu M. The Potential Role of IL-33 in Renal Transplant Recipients with Chronic Allograft Dysfunction. Annals of transplantation. 2016; 21:611–8.
- Legris T, Picard C, Todorova D, Lyonnet L, Laporte C, Dumoulin C, Nicolino-Brunet C, Daniel L, Loundou A, Morange S, Bataille S, Vacher-Coponat H, Moal V, et al. Antibody-Dependent NK Cell Activation Is Associated with Late Kidney Allograft Dysfunction and the Complement-Independent Alloreactive Potential of Donor-Specific Antibodies. Frontiers in immunology. 2016; 7:288.
- Dąbrowska-Żamojcin E, Dziedziejko V, Safranow K, Domański L, Słuczanowska-Głabowska S, Pawlik A. STAT4 gene polymorphism in patients after renal allograft transplantation. Central-European journal of immunology. 2016; 41:255–9.
- Cherukuri A, Rothstein DM, Clark B, Carter CR, Davison A, Hernandez-Fuentes M, Hewitt E, Salama AD, Baker RJ. Immunologic human renal allograft injury associates with an altered IL-10/TNF-alpha expression ratio in regulatory B cells. Journal of the American Society of Nephrology : JASN. 2014; 25:1575–85.
- Nogare AL, Dalpiaz T, Pedroso JA, Montenegro RM, Pegas KL, Veronese FV, Gonçalves LF, Manfro RC. Expression of fibrosis-related genes in human renal allografts with interstitial fibrosis and tubular atrophy. Journal of nephrology. 2013; 26:1179–87.
- 11. Hirt-Minkowski P, Marti HP, Hönger G, Grandgirard D, Leib SL, Amico P, Schaub S. Correlation of serum and urinary matrix metalloproteases/tissue inhibitors of metalloproteases with subclinical allograft fibrosis in renal transplantation. Transplant immunology. 2014; 30:1–6.
- Radwan-Oczko M, Boratynska M, Zietek M, Zoledziewska M, Jonkisz A. The relationship of transforming growth factorbeta1 gene polymorphism, its plasma level, and gingival overgrowth in renal transplant recipients receiving different immunosuppressive regimens. Journal of periodontology. 2006; 77:865–73.
- Dunning AM, Ellis PD, McBride S, Kirschenlohr HL, Healey CS, Kemp PR, Luben RN, Chang-Claude J, Mannermaa A, Kataja V, Pharoah PD, Easton DF, Ponder BA, et al. A transforming growth factorbetal signal peptide variant increases secretion *in vitro* and is associated with increased incidence of invasive breast cancer. Cancer research. 2003; 63:2610–5.

- Bidwell J, Keen L, Gallagher G, Kimberly R, Huizinga T, McDermott MF, Oksenberg J, McNicholl J, Pociot F, Hardt C, D'Alfonso S. Cytokine gene polymorphism in human disease: on-line databases, supplement 1. Genes and immunity. 2001; 2:61–70.
- 15. Jiménez-Sousa MA, Fernández-Rodríguez A, Heredia M, Tamayo E, Guzmán-Fulgencio M, Lajo C, López E, Gómez-Herreras JI, Bustamante J, Bermejo-Martín JF, Resino S. Genetic polymorphisms located in TGFB1, AGTR1, and VEGFA genes are associated to chronic renal allograft dysfunction. Cytokine. 2012; 58:321–6.
- Awad MR, El-Gamel A, Hasleton P, Turner DM, Sinnott PJ, Hutchinson IV. Genotypic variation in the transforming growth factor-beta1 gene: association with transforming growth factor-beta1 production, fibrotic lung disease, and graft fibrosis after lung transplantation. Transplantation. 1998; 66:1014–20.
- Nikolova PN, Ivanova MI, Mihailova SM, Myhailova AP, Baltadjieva DN, Simeonov PL, Paskalev EK, Naumova EJ. Cytokine gene polymorphism in kidney transplantation impact of TGF-beta 1, TNF-alpha and IL-6 on graft outcome. Transplant immunology. 2008; 18:344–8.
- Rechenberg DK, Galicia JC, Peters OA. Biological Markers for Pulpal Inflammation: A Systematic Review. PloS one. 2016; 11:e0167289.
- Biondi-Zoccai G, Lotrionte M, Landoni G, Modena MG. The rough guide to systematic reviews and meta-analyses. HSR Proc Intensive Care Cardiovasc Anesth. 2011; 3:161–73.
- 20. Park JY, Park MH, Park H, Ha J, Kim SJ, Ahn C. TNF-alpha and TGF-beta1 gene polymorphisms and renal allograft rejection in Koreans. Tissue antigens. 2004; 64:660–6.
- 21. Tan TLJ. Relationship between cytokine gene polymorphisms and chronic rejection in kidney allograft recipients and donors. Chin J Organ Transplant. 2007; 28:617–20.

- Li CJ, Yu LX, Xu J, Fu SJ, Deng WF, DU CF, Wang YB. [Association between transforming growth factor beta-1 gene polymorphism and chronic allograft nephropathy]. [Article in Chinese]. Nan Fang Yi Ke Da Xue Xue Bao. 2007; 27:535–7.
- 23. YLTLW W. Relationship between cytokine genotype and cyslosporine A in chronic renal allograft rejection. Strait Pharmaceutical Journal. 2007; 19:75–7.
- Melk A, Henne T, Kollmar T, Strehlau J, Latta K, Offner G, Jhangri GS, Ehrich JH, Von Schnakenburg C. Cytokine single nucleotide polymorphisms and intrarenal gene expression in chronic allograft nephropathy in children. Kidney international. 2003; 64:314–20.
- Gendzekhadze K, Rivas-Vetencourt P, Montano RF. Risk of adverse post-transplant events after kidney allograft transplantation as predicted by CTLA-4 +49 and TNF-alpha -308 single nucleotide polymorphisms: a preliminary study. Transplant immunology. 2006; 16:194–9.
- Ruhl S, Hamberger S, Betz R, Sukkar T, Schmalz G, Seymour RA, Hiller KA, Thomason JM. Salivary proteins and cytokines in drug-induced gingival overgrowth. Journal of dental research. 2004; 83:322–6.
- Masola V, Granata S, Bellin G, Gambaro G, Onisto M, Rugiu C, Lupo A, Zaza G. Specific heparanase inhibition reverses glucose-induced mesothelial-to-mesenchymal transition. Nephrology, dialysis, transplantation. 2017.
- Cambien F, Ricard S, Troesch A, Mallet C, Générénaz L, Evans A, Arveiler D, Luc G, Ruidavets JB, Poirier O. Polymorphisms of the transforming growth factor-beta 1 gene in relation to myocardial infarction and blood pressure. The Etude Cas-Temoin de l'Infarctus du Myocarde (ECTIM) Study. Hypertension. 1996; 28:881–7.