Harms and benefits of adoptive immunotherapy for postoperative hepatocellular carcinoma: an updated review

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ABSTRACT

The harms and benefits of adoptive immunotherapy (AIT) for patients with postoperative hepatocellular carcinoma (HCC) are controversial among studies. This study aims to update the current evidence on efficacy and safety of AIT for patients with HCC who have received curative therapy. Electronic databases were systematically searched to identify randomized controlled trials (RCTs) and cohort studies evaluating adjuvant AIT for patients with HCC after curative therapies. Recurrence and mortality were compared between patients with or without adjuvant AIT. Eight RCTs and two cohort studies involving 2,120 patients met the eligibility criteria and were meta-analyzed. Adjuvant AIT was associated with significantly lower recurrence rate than curative therapies alone at 1 year [risk ratio (RR) 0.64, 95%CI 0.49-0.82], 3 years (RR 0.85, 95%CI 0.79-0.91) and 5 years (RR 0.90, 95%CI 0.85-0.95). Similarly, adjuvant AIT was associated with significantly lower mortality at 1 year (RR 0.64, 95%CI 0.52-0.79), 3 years (RR 0.73, 95%CI 0.65-0.81) and 5 years (RR 0.86, 95%CI 0.79-0.94). Short-term outcomes were confirmed in sensitivity analyses based on RCTs or choice of a fixed- or random-effect meta-analysis model. None of the included patients experienced grade 3 or 4 adverse events. Therefore, this update reinforces the evidence that adjuvant AIT after curative treatment for HCC lowers risk of recurrence and mortality.

INTRODUCTION

Official guidelines [1, 2] identify hepatic resection and radiofrequency ablation (RFA) as two mainstay curative treatments for very early or early hepatocellular carcinoma (HCC). However, 5-year disease-free survival (DFS) associated with these treatments is only about 37% [3, 4]. Guidelines recommend transarterial chemoembolization (TACE) for intermediate or advanced HCC [1, 2], but progression-free survival (PFS) is also unsatisfactory [5, 6]: even after more aggressive hepatic resection, the 5-year recurrence rate can be as high as 74% [7, 8]. These data indicate that even after curative surgery, patients with HCC have poor prognosis, highlighting the need for effective adjuvant therapies that improve patient outcomes.

Many postoperative or adjuvant therapies have been described for improving the prognosis of patients with HCC, including adjuvant adoptive immunotherapy (AIT) [9,10]. Two systematic reviews from 2012 concluded that adjuvant AIT for patients with HCC after curative therapies may reduce recurrence rate but may not improve overall survival (OS) [11, 12]. These reviews included only a few small studies [13-18]. Since then, additional randomized controlled trials (RCTs) [19, 20] and cohort studies [21, 22] have been published with inconsistent findings. Therefore we wanted to perform an updated meta-analysis of the literature to gain a comprehensive

Table 1: Baseline characteristics o	f included studies.
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Study	Country	Study design	Surgery method	Child-Pugh score A/B, n (%)	Cirrhosis, <i>n</i> (%)	HBV/HCV, <i>n</i> (%)
Dong et al. 2009	China	RCT	Curative resection	102/25	101	96/NR
Huang et al. 2013	China	Retrospective	TACE+RFA	150/24 (86/14)	66	135/NR
Kawata et al. 1995	Japan	RCT	Curative resection + adriamycin	NR	14	NR
Lee, et al. 2015	Korea	RCT	Curative resection, RFA, or PEI	226/0	146	192/23
Pan et al. 2015	China	Retrospective	Curative resection	NR	NR	866/NR
Takayama et al. 2000	Japan	RCT	Curative resection	104/46	73	29/99
Weng et al. 2008	China	RCT	TACE+ RFA	69/16	NR	NR/NR
Xie et al. 2000	China	RCT	Curative resection + TACE	NR	NR	NR
Xu et al. 2016	China	RCT	Curative resection	200/0	113	171/NR
Zhou et al. 1995	China	RCT	Curative resection	NR	NR	NR/NR

Abbreviations: AIT, adoptive immunotherapy; HBV, hepatitis B virus; HCV, hepatitis C virus; NR, not reported; PEI, percutaneous ethanol injection; RCT, randomized controlled trial; RFA, radiofrequency ablation; TACE, transarterial chemoembolization.

understanding of the available evidence on the safety and efficacy of adjuvant AIT.

RESULTS

Description of studies

A total of 538 studies were identified, which decreased to 256 after duplicates were removed. Screening the titles and abstracts led to a final set of 20 studies that were read in full [13-22, 23-32]. Of these, six studies [23-28] were excluded because they contained subsets of patients already contained in larger studies [14, 15, 21, 22]. Three studies investigating AIT for patients with advanced HCC were excluded [29-31], and another study investigating a different type of postoperative immunotherapy was excluded [32]. In the end, 8 RCTs [13-20] and 2 cohort studies [21, 22] involving 1,079 AIT-treated and 1041 untreated patients were included in the meta-analysis (Figure 1, Table 1).

All studies were on patient populations in Asia. Patients in two studies had undergone a sequence of TACE followed by RFA prior to AIT [16, 21]. Patients in all other studies had undergone hepatic resection prior to AIT [13-15, 17-20, 22], with patients in one trial also undergoing postoperative transarterial adriamycin chemotherapy [14], and patients in another also receiving postoperative TACE [17]. One trial [13] contained two AIT-treated arms, one treated with 3 cycles and the other with 6 cycles. Data from the two arms were combined. Of all patients in the trial by Zhou et al. [18], only those who underwent resection alone or resection followed by adjuvant AIT were included in the present meta-analysis; this trial reported recurrence data out to 1 year only [18]. Across all studies in the meta-analysis, follow-up ranged from 18 months [16] to more than 6.5 years [21] (Table 2).

The present update substantially expands on the two previous systematic reviews comparing recurrence and mortality in patients receiving adjuvant AIT following curative therapies [11, 12]. The present work contains two RCTs [19, 20] and two cohort studies [21, 22], involving 1631 patients, that were not included in those previous reports.

Quality of the included studies

Risks of bias in the studies in this meta-analysis were detailed in Table 3. The methodological quality was high in two studies [19, 20] (accounting for 20% of the total patient population), moderate in two [13, 15] (accounting for 13% of total patients) and low in the remaining six [14, 16, 17, 18, 21, 22] (accounting for 67% of total patients).

Efficacy

Safety and efficacy data reported by each of the 10 studies in this meta-analysis [13-22] were summarized

in Table 2. Eight of the 10 studies reported that adjuvant AIT significantly improved DFS or PFS (all P < 0.05) [13, 15-19, 21, 22], while one small RCT [17] and two retrospective studies [21, 22] reported that adjuvant AIT significantly improved OS (all P < 0.05).

Meta-analysis of all 10 studies [13-22] suggested that adjuvant AIT was associated with significantly lower recurrence rate than curative therapies alone at 1 year (RR 0.64, 95%CI 0.49-0.82), 2 years (RR 0.70, 95%CI 0.59-0.84), 3 years (RR 0.85, 95%CI 0.79-0.91), and 5 years (RR 0.90, 95%CI 0.85-0.95) (Figure 2). Similar results were obtained using a random- or fixed-effects metaanalysis model. After excluding the two retrospective studies [21, 22], meta-analysis of the remaining 483 AIT-treated patients and 432 controls confirmed the recurrence benefit of adjuvant AIT at 1 year (RR 0.54, 95%CI 0.42-0.71), 2 years (RR 0.63, 95%CI 0.52-0.76) and 3 years (RR 0.81, 95%CI 0.71-0.93) (all P < 0.05). However, adjuvant AIT did not significantly reduce 5-year recurrence rate in this sensitivity analysis (RR 0.92, 95%CI 0.83-1.02).

Meta-analysis of 8 studies [13-15, 17, 19-22] suggested that adjuvant AIT was associated with significantly lower mortality than curative therapies alone at 1 year (RR 0.64, 95%CI 0.52-0.79), 2 years (RR 0.72, 95%CI 0.63-0.83), 3 years (RR 0.73, 95%CI 0.65-0.81),

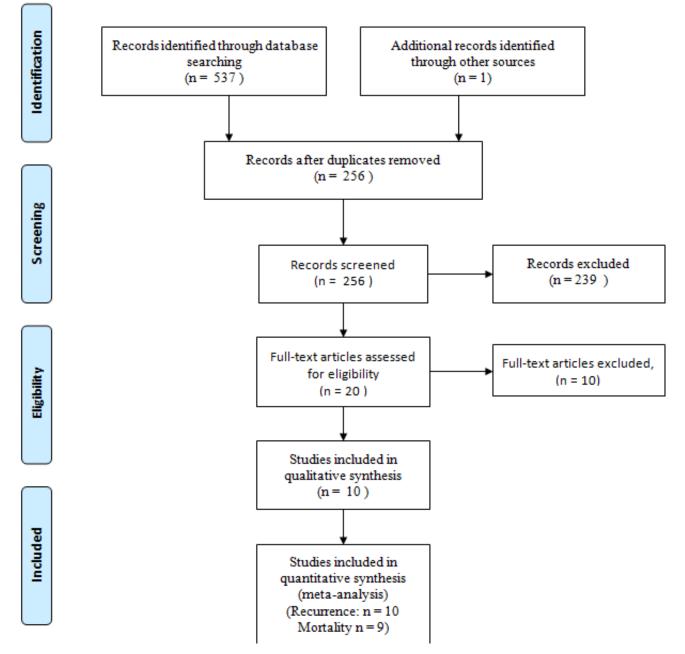


Figure 1: Flow chart of study selection.

	Recruitment	Sample	During and dama	Follow up	differenc		
Study	(T/C)	DFS or PFS	OS	Adverse events			
Dong et al. 2009	2000-2002	84/43	Group I: 3 cycles of CIK (1.0- 2.0×1010); Group II: 6 cycles of CIK (1.0- 2.0×1010)	>5 yr	DFS, P = 0.001 or $0.004*$	OS, <i>P</i> = 0.884	No long-term events
Huang et al. 2013	1999-2012	85/89	NR	Median, 6.5 yr (range, 0.4-14)	PFS, <i>P</i> = 0.001	OS, <i>P</i> = 0.001	No grade 3 or 4 adverse events
Kawata et al. 1995	1989-1990	12/12	13 mg/m2 adriamycin, IL-2, and 2.5x105 LAK daily for 3 weeks	NR	DFS, <i>P</i> = 0.182	OS, <i>P</i> = 0.936	No treatment- related deaths
Lee, et al. 2015	2008-2012	114/112	16 cycles of CIK cell agent	About 3 yr	DFS, $P = 0.01$	OS, P = 0.080	No grade 3 or 4 adverse events
Pan et al. 2015	2001-2009	511/520	At least 4 cycles CIK cells (1.0- 1.5x1010) via intravenous infusion	NR	PFS, <i>P</i> = 0.001	OS, <i>P</i> = 0.014	NR
Takayama et al. 2000	1992-1995	76/74	5 cycles of lymphocytes (IL-2 + Anti-CD3) (7.1×1010)	Median, 4.4 yr (range, 0.2-6.7)	DFS, $P = 0.010$	OS, <i>P</i> = 0.090	No grade 3 or 4 adverse events
Weng et al. 2008	2002-2004	45/40	$\begin{array}{cccc} 39 & \text{patients} \\ \text{received 8 cycles} \\ \text{of CIK} & (1.0-1.5\times1010); & 6 \\ \text{patients} & \text{received} \\ 10 & \text{cycles of CIK} \\ (1.0-1.5\times1010) \end{array}$	Median, 1.5 yr	DFS, $P = 0.012$	100% vs. 100%	No grade 3 or 4 adverse events
Xie et al. 2000	1994-1996	21/21	TACE + transarterial injection 1x109 LAK/ IL-2 (1x106 U)	NR	DFS, <i>P</i> < 0.05	OS, <i>P</i> < 0.05	NR
Xu et al. 2016	2008-2013	100/100	4 cycles CIK cells (1.0-1.5x1010) via intravenous infusion	Median, 3.2 (range, 0.3-6.1) years	DFS, $P = 0.334$	OS, <i>P</i> = 0.141	No grade 3 or 4 adverse events
Zhou et al. 1995	1992-1992	31/30	4 cycles of LAK + IL-2	NR	DFS, <i>P</i> < 0.05	NR	NR

Table 2: Study-level outcomes for HCC patients receiving adjuvant adoptive immunotherapy after curative therapies.

Abbreviations: AIT, adoptive immunotherapy; CIK, cytokine-induced killer cells; DFS, disease-free survival; IL-2, interleukin-2; LAK, lymphokine-activated killer cells; NR, not reported; OS, overall survival rate; PFS, progression-free survival; TACE, transarterial chemoembolization.

* Group I or II compared to control group.

and 5 years (RR 0.86, 95%CI 0.79-0.94) (all P < 0.05; Figure 3). Similar results were obtained using a randomor fixed-effects meta-analysis model. Sensitivity analysis using data from only the 6 RCTs [13-15, 17, 19, 20] supported a benefit of adjuvant AIT for mortality at 1 year (RR 0.39, 95%CI 0.21-0.72) and 2 years (RR 0.51, 95%CI 0.34-0.76), 3 years (RR 0.71, 95%CI 0.55-0.92), but not at 5 years (RR 0.99, 95%CI 0.83-1.19).

AIT-related adverse events

None of the 10 studies in the meta-analysis reported hospital deaths or serious adverse events attributed to adjuvant AIT. The most frequent adverse events due to

Study or Subgroup	AIT Events		Contr		Weight	Risk Ratio M-H, Fixed, 95% Cl	Risk Ratio M-H, Fixed, 95% Cl
3.1.1 1-year	Lycing	Total	240103	Total	recigint	in the trade of th	
Dong 2009	14	84	7	43	2.1%	1.02 [0.45, 2.35]	
Huang 2013	38	85	58	89	12.7%	0.69 [0.52, 0.91]	
Kawata 1995	2	12	3	12	0.7%	0.67 [0.13, 3.30]	
Lee 2015	15	114	26	112	5.9%	0.57 [0.32, 1.01]	
Pan 2015	235	511	263	520	58.5%	0.91 [0.80, 1.03]	-
Takayama 2000	13	76	30	74	6.8%	0.42 [0.24, 0.74]	
Weng 2008	4	45	12	40	2.9%	0.30 [0.10, 0.85]	
Xie 2000	O	21	6	21	1.5%	0.08 [0.00, 1.28]	←─────┼─
Xu 2016	14	100	23	100	5.2%	0.61 [0.33, 1.11]	
Zhou 1995	10	31	17	30	3.9%	0.57 [0.31, 1.04]	
Subtotal (95% CI)		1079			100.0%	0.77 [0.69, 0.86]	•
Total events	345		445				
Heterogeneity: Chi ² =	20.48, df	= 9 (P =	= 0.02); l ²	= 56%)		
Test for overall effect:							
3.1.2 2-year							
Dong 2009	28	84	30	43	6.9%	0.48 [0.33, 0.69]	_ _
Huang 2013	55	85	78	89	13.3%	0.74 [0.62, 0.88]	
Kawata 1995	6	12	,0	12	1.6%	0.67 [0.35, 1.28]	
Lee 2015	23	114	36	112	6.3%	0.63 [0.40, 0.99]	_ _
Pan 2015	291	511	333	520	57.4%	0.89 [0.81, 0.98]	
Takayama 2000	28	76	41	74	7.2%	0.66 [0.46, 0.95]	_
Xie 2000	4	21	10	21	1.7%	0.40 [0.15, 1.08]	
Xu 2016	27	100	32	100	5.6%	0.84 [0.55, 1.30]	
Subtotal (95% CI)		1003	02		100.0%	0.79 [0.73, 0.86]	•
Total events	462		569				
Heterogeneity: Chi ² =		= 7 (P =	= 0.01); I ^z	= 60%	,		
Test for overall effect:							
3.1.3 3-year							
Dong 2009	58	84	34	43	6.9%	0.87 [0.71, 1.08]	
Huang 2013	60	85	83	89	12.5%	0.76 [0.65, 0.88]	
Kawata 1995	6	12	9	12	1.4%	0.67 [0.35, 1.28]	
Lee 2015	40	114	49	112	7.6%	0.80 [0.58, 1.11]	_ _
Pan 2015	318	511	364	520	55.4%	0.89 [0.81, 0.97]	
Takayama 2000	40	76	51	74	7.9%	0.76 [0.59, 0.99]	
Xie 2000	12	21	18	21	2.8%	0.67 [0.44, 1.00]	<u>_</u>
Xu 2016	34	100	36	100	5.5%	0.94 [0.65, 1.38]	-
Subtotal (95% CI)		1003			100.0%	0.85 [0.79, 0.91]	♦
Total events	568		644				
Heterogeneity: Chi ² =		7 (P =		:0%			
Test for overall effect:	Z= 4.72 ((P < 0.0	0001)				
3.1.4 5-year							
Dong 2009	66	84	38	43	8.7%	0.89 [0.76, 1.04]	
Huang 2013	69	85	85	89			-
Pan 2015	342	511	385	520			
Takayama 2000	62	76	64	74		0.94 [0.82, 1.08]	7
Subtotal (95% CI)	02	756	70		100.0%		•
Total events	539		572				
Heterogeneity: Chi ² =	1.46, df=	3 (P =	0.69); l² =	:0%			
Test for overall effect:	Z = 3.55 ((P = 0.0	1004)				
							0.1 0.2 0.5 1 2 5 1 Favours AIT Favours control

Figure 2: Recurrence rate of meta-analysis comparing the efficacy of adjuvant adoptive immunotherapy (AIT) with curative treatment alone.

Study	Random allocation (description of procedure)	Concealment of random allocation	Blinding of persons who assess treatment effects	Intention-to-treat analysis
Dong et al. 2009	+	-	-	-
Huang et al. 2013	-	-	-	-
Kawata et al. 1995	-	-	-	-
Lee, et al. 2015	+	+	-	+
Pan et al. 2015	-	-	-	-
Takayama et al. 2000	+	-	-	+
Weng et al. 2008	-	-	-	-
Xie et al. 2000	-	-	-	-
Xu et al. 2016	+	+	+	+
Zhou et al. 1995	-	-	-	-

Table 3: Assessment of methodological quality (internal validity) of included studies.

AIT were grade 1 fever (defined as persistent or transient temperature of 37.5-39.3°C) and chills. The study with the highest frequency of persistent fever reported it in 5 of 84 (6.0%) patients [13], and none of the 5 was able to complete AIT per protocol because of this condition. In all patients experiencing fever in the meta-analysis, the condition was easily controlled with symptomatic therapies. Rare adverse events included headache, nausea, myalgia, fatigue, dizziness, itching, and tachycardia. All adverse events were grade 1 or 2 and self-limiting. In no case did adverse events cause patients to delay or stop treatment, except for the 5 patients with persistent fever mentioned above. No cases of infection, hepatic deterioration, pulmonary symptoms or autoimmune disorder were reported in the 10 studies.

Assessment of publication bias

Funnel plots of the 10 studies in the meta-analysis showed a symmetrical shape, suggesting minimal risk of publication bias (Figures 4 and 5).

DISCUSSION

HCC is associated with a high recurrence rate, even after curative treatment; in fact, recurrence is the primary cause of death of all patients with HCC. Even after hepatic resection of HCC, patients with large/multinodular HCC can show 5-year DFS of 26%, while this rate can be as low as 18% in those with macrovascular invasion, based on a systematic review of more than 14,000 patients [33]. For such patients, adjuvant TACE shows promise for reducing recurrence and mortality [34]. For patients with hepatitis B virus-related HCC, postoperative antiviral therapy can be safe and effective treatment [35, 36]. However, some HCC patients are unfit for TACE or antiviral therapy after surgery. For these patients, and for those with low immune function, which is associated with HCC recurrence [37], adjuvant AIT may prevent tumor relapse. Adjuvant AIT involves transferring immune effectors into the cancer patient in the hopes of stimulating specific anti-tumor immune responses [38]. Such stimulation may counterbalance the strongly immunosuppressive microenvironment in the liver [39].

The present meta-analysis updates two systematic reviews [11, 12] from 2012 examining the safety and efficacy of adjuvant AIT for HCC patients who have received curative therapies. In contrast to those previous reports [11, 12], this update provides strong evidence that AIT can significantly reduce the rate of tumor recurrence and mortality. The discrepancy between our findings and those of previous systematic reviews likely reflects the more than two decades spanned by the literature, with the first RCTs on AIT for postoperative HCC published in 1995 [14, 18] and the most recent in 2016 [20], combined with rapid scientific and technological advances in AIT [40, 41]. In addition, no international guidelines or standards exist regarding route of AIT administration, dosing, or cycles. As a result, clinicians can vary substantially in what immune effector cells they use for AIT and what dosing/cycling protocols they follow. Indeed, in the present meta-analysis, AIT was based on three types of immunological effector cells: anti-CD3-activated peripheral blood lymphocytes, cytokineinduced killer cells, and lymphokine-activated killer cells. AIT was administered via injection into the intrahepatic artery [17, 18] or via intravenous infusion [15, 16, 19-22]. The number of cycles varied from one [17] to 16 [19]. Such heterogeneity highlights the importance of evidence updates like the present one, and the need for systematic assessment and optimization of AIT protocols, perhaps even tailored to HCC type or treatment history.

Our meta-analysis of RCT data suggests that adjuvant AIT can significantly reduce recurrence and mortality at 1, 2 and 3 years, but not 5 years. This may mean that AIT-mediated immune boosting can eliminate small intrahepatic metastases, but it does not prevent multicentric relapse in remnant liver. This hypothesis is consistent with the findings of one study [20] in our metaanalysis that reported that the ability of adjuvant AIT to prevent tumor recurrence was more obvious in the short term and less so in the long term, and that its ability to prolong time to recurrence was greater in patients with tumors >5 cm, moderately or poorly differentiated tumors, or preoperative α -fetoprotein levels \geq 25 ng/mL. Though

Study or Subgroup	Experim Events		Contr Events		Weight	Risk Ratio M-H, Fixed, 95% Cl	Risk Ratio M-H, Fixed, 95% Cl
4.1.1 1-year					0	, , , ,	
Dong 2009	11	84	6	43	4.3%	0.94 [0.37, 2.37]	
Huang 2013	14	85	19	89	10.1%	0.77 [0.41, 1.44]	_ _
Kawata 1995	0	12	2	12	1.4%	0.20 [0.01, 3.77]	
Lee 2015	3	114	5	112	2.7%	0.59 [0.14, 2.41]	
Pan 2015	90	511	133	520	71.5%	0.69 [0.54, 0.87]	
Takayama 2000	1	76	4	74	2.2%	0.24 [0.03, 2.13]	
Xie 2000	0	21	4	21	2.4%	0.11 [0.01, 1.94]	·
<u 2016<="" td=""><td>1</td><td>100</td><td>10</td><td>100</td><td>5.4%</td><td>0.10 [0.01, 0.77]</td><td></td></u>	1	100	10	100	5.4%	0.10 [0.01, 0.77]	
Subtotal (95% CI)		1003		971	100.0%	0.64 [0.52, 0.79]	•
Fotal events	120		183				
Heterogeneity: Chi ² =	= 7.34, df =	7 (P = 0	.39); I ² = 6	5%			
Fest for overall effect	: Z= 4.16 (I	P < 0.00	01)				
4.1.2 2-year							
Dong 2009	13	84	7	43	2.8%	0.95 [0.41, 2.21]	
Huang 2013	24	85	35	89	10.5%	0.72 [0.47, 1.10]	+
Kawata 1995	1	12	2	12	0.6%	0.50 [0.05, 4.81]	
_ee 2015	5	114	13	113	4.0%	0.38 [0.14, 1.03]	
Pan 2015	179	511	234	520	71.0%	0.78 [0.67, 0.91]	
Takayama 2000	6	76	11	74	3.4%	0.53 [0.21, 1.36]	
<ie 2000<="" td=""><td>3</td><td>21</td><td>9</td><td>21</td><td>2.8%</td><td>0.33 [0.10, 1.06]</td><td></td></ie>	3	21	9	21	2.8%	0.33 [0.10, 1.06]	
(u 2016	7	100	16	100	4.9%	0.44 [0.19, 1.02]	
Subtotal (95% CI)		1003		972	100.0%	0.72 [0.63, 0.83]	•
Total events	238		327				
Heterogeneity: Chi² = Test for overall effect)%			
		0.00	001)				
4.1.3 3-year							
Dong 2009	29	84	15	43	4.6%	0.99 [0.60, 1.64]	
Huang 2013	31	85	50	89	11.4%	0.65 [0.46, 0.91]	
Kawata 1995	3	12	3	12	0.7%	1.00 [0.25, 4.00]	
_ee 2015	19	114	25	112	5.9%	0.75 [0.44, 1.28]	
Pan 2015	204	511	278	520	64.6%	0.75 [0.65, 0.85]	• • • • • • • • • • • • • • • • • • •
Fakayama 2000	10	76	19	74	4.5%	0.51 [0.26, 1.03]	
<ie 2000<="" td=""><td>8</td><td>21</td><td>15</td><td>21</td><td>3.5%</td><td>0.53 [0.29, 0.98]</td><td></td></ie>	8	21	15	21	3.5%	0.53 [0.29, 0.98]	
Ku 2016 Subtatal (OEV, CD	13	100	20	100	4.7%	0.65 [0.34, 1.23]	
Subtotal (95% CI)		1003	405	971	100.0%	0.73 [0.65, 0.81]	•
Fotal events	317		425				
Heterogeneity: Chi² = Test for overall effect		-		1%			
4.1.4 5-year			-				
	50		27	40	7.00	0.00 (0.74, 4.04)	\perp
Dong 2009	52	84 95	27	43	7.6%	0.99 [0.74, 1.31]	
Huang 2013 Pop 2015	47	85 514	66 222	89 520	13.8%	0.75 [0.59, 0.94]	
Pan 2015 Fakayama 2000	268	511	322	520	68.2%	0.85 [0.76, 0.94]	
Takayama 2000 Subtotal (95% CI)	49	76 756	48	74 726	10.4% 100.0%	0.99 [0.78, 1.26] 0.86 [0.79, 0.94]	
Total events	416		463	. 20			
Heterogeneity: Chi ² =		3 (P = 0		23%			
Test for overall effect							
							· · · · · · · · · · · · · · · · · · ·
est for subaroup dif	fforoncoc [.] (Chi² = 10	1 0 8 df - 1	3 (P = 1	0.01) 🖻 –	77.7%	Favours [experimental] Favours [control]

Test for subgroup differences: $Chi^2 = 10.98$, df = 3 (P = 0.01), $l^2 = 72.7\%$

Figure 3: Mortality of meta-analysis comparing the efficacy of adjuvant adoptive immunotherapy (AIT) with curative treatment alone.

our results were supported by previous systematic reviews [42, 43], the effects of adjuvant AIT on HCC recurrence in the short and long term should be investigated in greater detail.

Patients in two of the included studies underwent minimally invasive treatments [16, 21]. In one of these studies, no patient died during the 1.5-year follow-up [16]. In the other study [21], adjuvant AIT significantly improved PFS and OS. These results raise the possibility that the combination of minimally invasive treatments and postoperative AIT may exert synergistic effects. For example, since RFA is known to stimulate the differentiation of natural killer cells and boost their activity [44], it is possible that combining RFA with AIT may further boost immune function and reduce the rate of tumor recurrence. Future studies should investigate the possibility that for suitable patients, the combination of RFA and AIT may be superior to the combination of resection and AIT.

AIT may be based either on human leukocyte antigen-restricted or unrestricted strategies [45]. Cytokineinduced killer cells and lymphokine activated killer cells are heterogeneous mixture of immune effector cells that feature a mixed T- and natural killer cell-like phenotype in their terminally-differentiated CD3+CD56+ subset. Cytokine-induced killer cells and lymphokine activated killer cells can exhibit histocompatibility complexunrestricted cytotoxicity against a broad range of tumors [46]. Transferred T-cell-based cytotoxicity is the most probable mechanism for anti-CD3-activated peripheral blood lymphocytes [15]. These natural effectors carry out their antitumoral activities without identify and recognize the presence of specific tumor associated antigens expressed on the cells surface. The easy availability, high proliferation rate and widely major histocompatibility complex-unrestricted antitumor activity of three types of cells contribute to their particularly advantageous profile, making them an attractive approach for AIT. Micrometastatic HCC cells are plausible targets. Use of peripheral blood as the source of effectors is supported by the fact that tumour-specific cytotoxic T-cells can be isolated from the peripheral repertoire. However, the extent to which specific T-cell responses contribute to the best clinical outcome needs further clinical trials.

In the past two decades, the scientific interest is focused on oral multikinase inhibitor drugs. However, adjuvant oral multikinase inhibitors provided negative efficacy for HCC after surgery, RFA, or TACE [47-51], giving the space to explore new effective adjuvant therapies.

The findings of this meta-analysis that adjuvant AIT significantly reduces recurrence and mortality for postoperative HCC must be interpreted with caution. Surgical method, type of cytokines, number of infusion

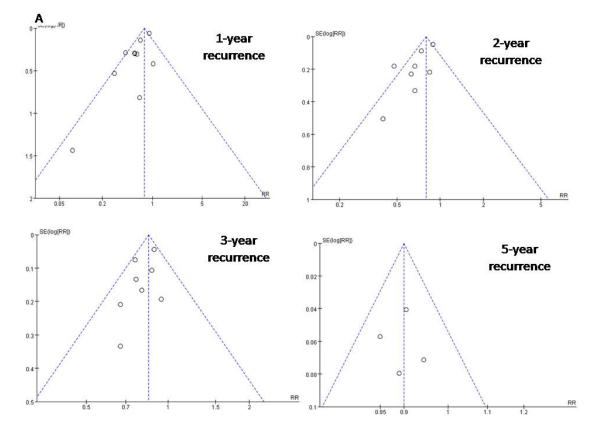


Figure 4: Funnel plots to detect any publication bias about recurrence rate.

cycles, and duration of maintenance AIT therapy varied among the included studies, creating substantial heterogeneity for which we could not control using sensitivity analyses. Therefore we displayed the efficacy results for each study individually in Table 2. In addition, length of follow-up varied across the studies and in some cases was too short to observe long-term efficacy of adjuvant AIT. As a result, meta-analysis of outcomes at 3 and 5 years had to be conducted on subsets of all included studies. Some studies did not clearly report procedures for randomization or allocation concealment, increasing the risk of selection or reporting bias. The fourth problem with this meta-analysis is that there are limitations in the original data, which are beyond our control, but nevertheless compromise the value of the study. We know very little about surveillance/screening methodology, diagnostic criteria for HCC, and stage systems for HCC in this meta-analysis. So, a large variability of post-treatment surveillance programs and diagnostic criteria among studies could be expected. The last relevant issue of this meta-analysis is the potential lack of external validity of the results for different populations and settings. All the included studies were conducted on patient populations in Asia. So, a high rate of hepatitis B virus infected patients with or without cirrhosis could be expected. This

population may be different in terms of clinical features and comorbidities from most cases of hepatitis C virusrelated or post- non-alcoholic steatohepatitis HCC from US and Europe.

Despite these limitations, our meta-analysis provides an updated picture of the evidence based on adjuvant AIT: AIT may be superior to either hepatic resection alone or the combination of TACE followed by RFA for postoperative HCC patients. The findings of the present meta-analysis should be verified and extended in further large trials with adequate follow-up. These studies should aim to expand the range of relevant endpoints examined, such as quality of life, duration of hospital stay, and cost-effectiveness. These studies should also examine the possible clinical benefits of multi-modal immune therapies.

MATERIALS AND METHODS

Literature search strategy

The most recent on-line versions of the following research databases were searched in June 2016 without

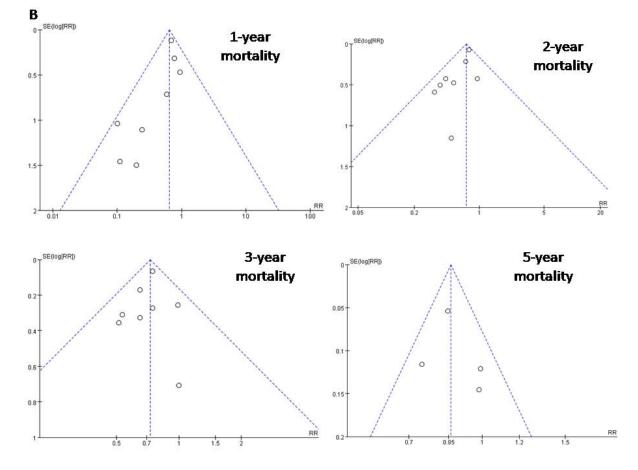


Figure 5: Funnel plots to detect any publication bias about mortality.

Cochrane Library (http:// language restrictions: onlinelibrary.wiley.com/cochranelibrary/search), Wiley Online Library, Science Direct, Web of Science, Chinese National Knowledge Infrastructure, Embase, and PubMed. The following search terms were used to identify comparative studies: 'hepatocellular carcinoma' or HCC or 'hepatic cancer' or 'hepatic tumor' or 'liver tumor' or 'liver cancer', and 'hepatic resection' or hepatectomy or 'liver resection' or 'transarterial chemoembolization' or 'radiofrequency ablation' or 'invasive treatment' or 'percutaneous ethanol infection', and 'adoptive immunotherapy' or 'cytokine induced killer cells' or 'lymphokine activated killer cells' or 'tumor infiltrating lymphocytes' or 'interleukin-2'. To avoid missing relevant studies, no filter was imposed to exclude non-controlled studies or non-RCTs. Relevant references were also searched manually to identify additional studies.

Inclusion criteria

We included in the meta-analysis full-length research studies that satisfied the following criteria: (a) the study compared the efficacy of curative therapies with or without adjuvant AIT for patients with HCC; (b) the study had a randomized control or cohort design; (c) all tumors were treated by curative procedures before AIT; (d) patients in the AIT and no-AIT arms received otherwise similar treatments; and (e) the study reported sufficient data for estimating risk ratios (RRs) with 95% confidence intervals (95%CIs). Curative therapies for HCC included hepatic resection, RFA, percutaneous ethanol injection, and liver transplantation; these therapies were considered curative if no residual tumor was observed one month after the initial therapy. TACE was defined as palliative therapy for the purposes of this meta-analysis.

Studies were excluded if they evaluated the efficacy of AIT for patients with liver metastases or with recurrent, advanced, or unresectable HCC. Conference abstracts and other forms of summary publication were also excluded. In the case of multiple studies apparently based on the same population, we included only the study with the largest number of participants.

Study identification and data extraction

Studies identified in literature searches were independently screened by two authors (B.-H.Y, R.-H.L), with discrepancies arbitrated by a third author (J.-H.Z). Two authors (B.-H.Y, R.-H.L) independently extracted the following data from included studies using a predefined template: author details, country, study design, surgery method, liver disease, recruitment period, sample size, follow-up period, interventions (drugs, schedules and numbers of therapy sessions), outcomes (positive and negative findings), and methodological quality. A third author (J.-H.Z) checked the extracted data against the original studies. Survival data were taken directly from tables or the text whenever possible; if such data were presented only in graphs, they were extracted by manual interpolation [33]. *P* values associated with inter-group differences in PFS, DFS, or OS were extracted directly from survival curves, text, or tables wherever possible.

Outcome measures

The primary outcomes in this meta-analysis were recurrence rate and mortality. The secondary outcome was treatment-related adverse events, which included treatment-related withdrawals and discontinuations.

Quality assessment

This meta-analysis was conducted in accordance with the Quality of Reporting of Meta-analyses (QUOROM) statement [52]. Two authors (B.-H.Y, R.-H.L) independently evaluated all included RCTs based on method of randomization, allocation concealment, blinding of outcome assessors, and use of intention-totreat analysis. RCTs were considered to be of low quality if they reported none of the items, of moderate quality if they reported on fewer than three items and of high quality if they reported on three or four items [12, 52]. Quasirandomized studies and cohort studies were defined to be of low quality.

Missing data

Meta-analysis was performed on an intentionto-treat basis. To assess attrition bias, we calculated recurrence and mortality using a 'worst-case' approach in which patients with missing data were counted as treatment failures (recurrence or death). For patients with missing data, we 'carried forward' data from the most recent measurement.

Statistical analysis

Review Manager 5.3 (Cochrane Collaboration) was used to analyze data from included studies. Due to the high likelihood of recurrence and mortality, RRs with corresponding 95%CIs were calculated for dichotomous outcomes using the Mantel-Haenszel method. Point estimates of RR were considered statistically significant when P < 0.05. Meta-analysis was carried out using a random-effects model if substantial heterogeneity according to an I-squared threshold was found among included studies; otherwise, the analysis was carried out

using a fixed-effect model [53]. If the two models gave different results, we reported both results. Heterogeneity was assessed by calculating I². Homogeneity between studies was analyzed using the χ^2 test, with significance set at P > 0.1. Publication bias was assessed by visual inspection of Begg's funnel plots. Sensitivity analyses excluding cohort studies and choice of random- or fixed-effect meta-analysis model were performed.

CONFLICTS OF INTEREST

The authors have declared that no competing interests exist.

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